



# The Relative changes in the Steady States of Per-Capita Output of Some Typical “Middle Income Trap” Countries and China

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**Abstract:** Based on a sample of 114 countries and regions, this paper uses the econometric method to show that, in terms of steady state of per-capita output, the relative positions of Brazil, Mexico, Malaysia, Turkey and South Africa in the sample generally remained slightly lower than the average level of the sample in the 1970-2019 period; China's relative position in the sample was extremely low in the 1970s, then continued to rise rapidly and caught up with the overall level of the above five countries in 2010s. Thus, even in terms of steady state of per-capita output, the above five countries were still typical “middle income trap” countries in the 1970-2019 period while China was not, but China started to face the “middle income trap” in 2010s. Next, combining the theory of convergence with the practical data of the above 5 countries and China, this paper analyses the reasons, respectively, for their different changes in the relative position in the sample. Finally, from the perspective of steady state of per-capita output, this paper gives some suggestions, respectively, for the above five countries and China to cross the “middle income trap” in the future.

**Keywords:** “middle income trap” country; China; steady state of per-capita output; conditional convergence; social infrastructure

## 1. INTRODUCTION

Many scholars have, from different perspectives, studied the reasons why some developing countries fell into and have been staying in the “middle income

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trap”, and actually they have made many valuable research results. For example, Kam (2014) believed that the productivity growth in Malaysian manufacturing sector was low and the lack of innovative capabilities of the Malaysian manufacturers undermined the industrial upgrading prospects. Luiz (2016) argued that South Africa was unlikely to move beyond middle-income status unless there was a dramatic policy shift which could bring appropriate investments and better human capital plan in the country. Ada and Acarođlu (2016) expressed that Turkey could overcome the “middle income trap” if it attached enough importance for public spending on education for a better human capital growth. Foxley and Stallings (2016) stressed that Latin America needed greater institutional capacity to promote innovation which in turn increased productivity. Dabús, Tohmé, and Caraballo (2016) pointed out that once the world’s demand for primary commodities did not increase or even decrease, Latin American countries that rely heavily on international prices of these commodities would fall into the “middle income trap”. Paus (2019) argued that the current globalisation process had shifted the goal posts for middle-income countries and increased the urgency for Latin American countries to develop domestic innovation capabilities to improve their poor productivity performance. Topal (2020) indicated that economic and institutional reform requirements maintained their priority in the political agenda in most middle-income countries, especially in Latin American countries.

Based on the above findings, one could draw the following conclusion: after reaching the middle-income level, the above developing countries failed to realise the transformation of their economic development strategies and modes, which resulted in difficulties in their industrial upgrading and the lack of endogenous driving force for their economic growth. As a result, they have long been among the “middle income trap” countries. The above conclusion is certainly pertinent, but the work of these scholars can still be improved. The reason for that is the more convincing quantitative analysis (such as quantitative analysis of Econometrics) was obviously less in their research work, which directly affected the academic value of their research results. So the future research works should try to make up for this shortcoming.

This paper uses the econometric method to make a study on the “middle income trap” countries from the perspective of steady state of per-capita output, which results from the Solow model. As the Solow model shows, for a given period, an economy’s per-capita output always converges toward its steady state of per-capita output in that period. It can also be inferred that developed countries’ steady states of per-capita output are usually much higher than

developing countries'. In addition, due to existence of capital accumulation and technological innovation, most countries (including developing and developed ones) experience growth in their steady states of per-capita output over time. The explanations for the above two statements will be given in Section 2. To catch up with developed countries, developing countries need to achieve a relative growth in the steady state of per-capita output, so it is worthwhile to investigate the relative changes in the steady states of per-capita output of important developing countries in a broad set of countries. For doing that, this paper builds an important concept: *the relative steady state of per-capita output*. The detailed explanation for this concept will be given as well in Section 2.

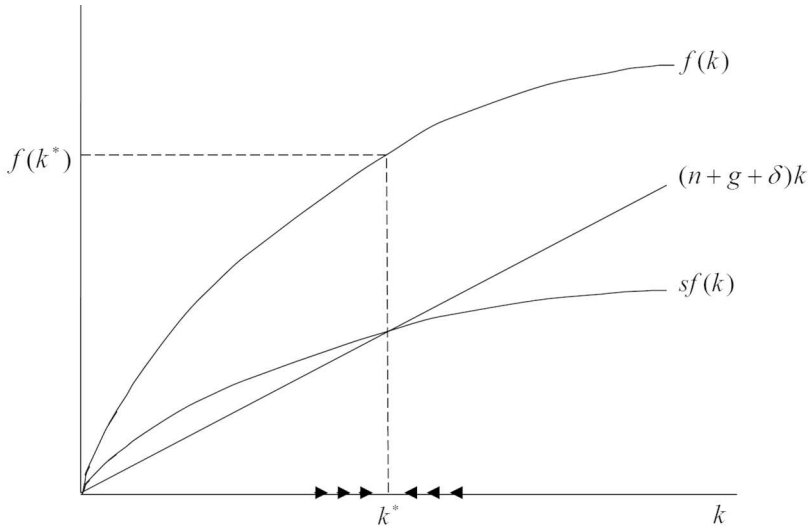
Through testing the hypothesis of conditional convergence, this paper obtained the estimates of the relative steady states of per-capita output of five typical "middle income trap" countries (Brazil, Mexico, Malaysia, Turkey and South Africa), China (the mainland of China, the same below) and the United States (as a representative of developed countries). These estimates were used to show the relative changes in the steady state of per-capita output of each of the above country in a test sample. Then, combining the theory of convergence with the practical data of the above countries, this paper made an analysis of the reasons for the relative changes in their steady states of per-capita output. Conclusions were given after a comparison of the five typical "middle income trap" countries and China.

The paper consists of seven sections. Section 1 is introduction. Section 2 is a brief review of previous studies on convergence. In Section 3, the regression equation to test the hypothesis of conditional convergence is described. In Section 4, the data and the empirical methodology used are described, and the details of results and analyses are also given. After Section 5 showing the paths of the relative steady states of per-capita output of the concerned countries, Section 6 provides an analysis of reasons for the relative changes in their steady states of per-capita output. Conclusions are given in Section 7.

## 2. A BRIEF REVIEW OF PREVIOUS STUDIES ON $\beta$ -CONVERGENCE

Most economists did their studies on convergence which stemmed from the Solow's classical growth model. The Solow model proposed the concept of steady state of per-capita output, and Figure 1 shows the details. In Figure 1, for an economy in a given period, the capital per unit of effective labour  $k$  converges toward its steady state  $k^*$ , so the output per unit of effective labour  $f(k)$  converges toward its steady state  $f(k^*)$ . Further, the output per unit of labour (i.e., per-

capita output)  $Af(k)$  converges toward its steady state  $Af(k^*)$ , where  $A$  denotes the effectiveness of labour in the given period.



**Figure 1: An economy's steady state in a given period**

$\beta$ -convergence exists for a set of economies. It is named after the speed of convergence  $\beta$  and consists of absolute convergence and conditional convergence. Absolute convergence means, the selected economies have similar steady state of per-capita output to converge. Conditional convergence means, the selected economies have different steady states of per-capita output to converge, respectively. Conditional convergence is obviously more common in a set of economies, so the previous studies on  $\beta$ -convergence generally focused on conditional convergence.

By the way, there are some other models which can be used for convergence study. For example, Phillips & Sul (2007) established a new model providing a new method to investigate convergence, which is regarded as an important contribution in the field of convergence and has actually been used frequently by many economists. Actually Phillips & Sul (2009) used their method to show that the growths of developed and developing countries would converge to different levels through displaying relevant transition parameters. But their method did not involve the steady state mentioned in the Solow model, so all previous studies using the method of Phillips & Sul did not give any information on the steady state mentioned in the Solow model.

This paper provides a study on convergence based on the steady state mentioned in the Solow model. It is well known, as for the steady state mentioned in Solow model, many economists have found the evidence of conditional convergence (e.g., Baumol (1986), Barro (1991), Mankiw, Romer, and Weil (1992), Caselli, Esquivel, and Lefort (1996), Lee, Pesaran, and Smith (1997), Panik and Rassekh (2002), Mathur (2005), McQuinn and Whelan (2007), Karras (2008), Cavenaile and Dubois (2011), Rath (2016), Stengos, Yazgan, and Ozkan (2018), etc), the main difference among their regression results focused on the estimate of the speed of convergence. But it is necessary to point out that their studies on convergence were made by using only one period rather than *several successive* sub-periods. The reason for that is most economists believed that the concept of convergence only applied to a long period, which consists of several decades or even hundreds of years. When talking about developing countries, they even argued that the lack of development was due to the distance to steady state, not the level of different steady states among countries. Their idea is still commonly recognised today, but this paper tries to make it possible to challenge their idea.

Firstly, the Solow model focuses on discussing the effects of some economic parameters on the steady state of per-capita output ( $Af(k^*)$ ). This model implies that, if the economic parameters ( $s$ ,  $n$ ) change or the effectiveness of labour ( $A$ ) changes between two different periods, a country may experience a change in its the steady state of per-capita output between the two periods. In reality, a country's the economic parameters and the effectiveness of labour will change at times, so the idea that a country's steady state remains unchanged during a long period is possibly wrong.

Furthermore, since the changes in steady state of per-capita output are usually different across countries, there will be a *relative change* in steady state of per-capita output of a country among a set of countries. Thus, one needs to find an indicator for such a relative change. Suppose there is a test sample which includes  $N$  countries, and let  $Y_i^*$  denote the steady state of per-capita output of country  $i$  for all  $i$  ( $i = 1, 2, 3, \dots, N$ ), so  $Y_i^* = A_i f(k_i^*)$  holds for country  $i$  for all  $i$ . Let  $\bar{Y}^*$  denote the average level of the  $N$  countries in the test sample, then  $Y_i^* / \bar{Y}^*$  denotes *the relative steady state of per-capita output* of country  $i$  for all  $i$ , that is the ratio of the steady state of per-capita output of country  $i$  for all  $i$  to the average level of all countries in the test sample. If a change in this ratio is significant, it means there is a relative change in the steady state of per-capita

output of country  $i$  for all  $i$  in the sample. In addition, the logarithmic version of this ratio,  $\log(Y_i^* / \bar{Y}^*)$ , can also be used to express the relative steady state of per-capita output of country  $i$  for all  $i$ , which is the operation of this paper.

Secondly, although the Solow model does not show whether a country is close to or obviously away from its steady state in a given period, it seems reasonable to believe a country is much possibly close to its steady state of per-capita output if the changes in its economic parameters ( $s$ ,  $n$ ) and effectiveness of labour ( $A$ ) are not significant for several decades. In addition, according to the Solow model, even supposing there are no significant differences in economic parameters between developing and developed countries, one can believe that most developing countries are lower than developed countries in the steady state of per-capita output, for their effectiveness of labour are much lower than the level of developed countries.

Thirdly, the convergence theory does imply that an economy's steady state of per-capita output exists for a given period, but this theory does not specify the length of the given period. Theoretically, an economy's steady state of per-capita output can exist in a relatively short period, such as a period of 10 years.

No matter whether a country is developing or developed, and whether it is close to or obviously away from its steady state, the country will always converge toward its steady state which may change over time, so it is surely worthwhile to investigate the relative changes in the steady state of a country among a broad set of countries. But the previously mentioned studies did not test the hypothesis of conditional convergence across successive sub-periods, so they did not assess whether there happened, across sub-periods, a significant change in a country's relative steady state of per-capita output. Such a change means a relative change in a country's steady state of per-capita output among a broad set of countries.

This paper undertakes such a study by testing the hypothesis of conditional convergence in a test sample of 114 countries and regions in 1970s, 1980s, 1990s, 2000s and 2010s. This paper also shows the paths of the relative steady states of per-capita output of five typical "middle income trap" countries (Brazil, Mexico, Malaysia, Turkey and South Africa), China and United States by using their estimates obtained in the above five successive sub-periods. A comparison of the seven paths provides some valuable information on the growths of the above seven countries.

### **3. THE REGRESSION EQUATION TO TEST THE HYPOTHESIS OF $\beta$ -CONVERGENCE**

The following equation was actually provided by Barro and Sala-I-Martin (2004). The only difference is that the equation on the page 466 of *Economic Growth*

(2004, 2nd ed.) shows the time interval  $T$  of observations is between year 0 and year  $T$ .

$$(1/T)\log(Y_{i,t}/Y_{i,t-T}) = \alpha_i - (1/T)(1 - e^{-\beta T})\log Y_{i,t-T} + u_{i,t}, \tag{1}$$

where the subscript  $t$  denotes year  $t$ ; the subscript  $i$  denotes economy  $i$ ;  $T$  denotes the time interval of observations between year  $t-T$  and year  $t$ ;  $Y_{i,t}$  denotes per-capita output of economy  $i$  for all  $i$  in year  $t$ , i.e.,  $Y_i = A_i f(k_i)$  holds for economy  $i$  for all  $i$ ;  $\beta$  denotes the average speed of convergence for all economies in a sample for a given period;  $\alpha_i = x_i + (1/T)(1 - e^{-\beta T})\log Y_i^*$ ,  $x_i$  denotes the technological progress rate of economy  $i$  for all  $i$  (i.e.,  $x_i = g_i$  holds for all  $i$ ), the natural number  $e \cong 2.718$ ,  $Y_i^*$  denotes the steady state of per-capita output of economy  $i$  for all  $i$  for a given period, so  $Y_i^* = A_i f(k_i^*)$  holds for economy  $i$  for all  $i$  for the period. The equation (1) implies the average annual growth rate (from year  $t-T$  to year  $t$ ) of per-capita output of economy  $i$  for all  $i$  depends positively on  $Y_i^*$  and negatively on  $Y_{i,t-T}$ .

In order to remove the time trend associated with the growth of technological progress ( $x_i$ ), Coulombe (2004) defined  $y_{i,t} = \log(Y_{i,t} / \bar{Y}_t)$ , where  $\bar{Y}_t$  is the cross section mean of  $Y_{i,t}$  in year  $t$  for all  $t$ . With this definition, the equation (4) can be obtained by transforming the equation (1); the details are shown as follows.

Firstly, the equation (1) can be rewritten as

$$(1/T)(\log Y_{i,t} - \log Y_{i,t-T}) = \alpha_i - (1/T)(1 - e^{-\beta T})\log Y_{i,t-T} + u_{i,t}, \tag{2}$$

Then take the mean over the number of economies  $N$  of this equation and obtain

$$(1/T)\left(\frac{1}{N}\sum_{i=1}^N \log Y_{i,t} - \frac{1}{N}\sum_{i=1}^N \log Y_{i,t-T}\right) = \frac{1}{N}\sum_{i=1}^N \alpha_i - (1/T)(1 - e^{-\beta T})\frac{1}{N}\sum_{i=1}^N \log Y_{i,t-T} + \frac{1}{N}\sum_{i=1}^N u_{i,t}$$

$$\text{or } (1/T)(\log \bar{Y}_t - \log \bar{Y}_{t-T}) = \bar{\alpha} - (1/T)(1 - e^{-\beta T})\log \bar{Y}_{t-T} + \bar{u}_t, \tag{3}$$

where  $\bar{Y}_t = \sqrt[N]{Y_{1,t} Y_{2,t} \dots Y_{N,t}}$ ;  $\bar{Y}_{t-T} = \sqrt[N]{Y_{1,t-T} Y_{2,t-T} \dots Y_{N,t-T}}$ ;  $\bar{\alpha} = \bar{x} + (1/T)(1 - e^{-\beta T})\log \bar{Y}^*$ ,

$\bar{x} = (1/N)\sum_{i=1}^N x_i$  and  $\bar{Y}^* = \sqrt[N]{Y_1^* Y_2^* \dots Y_N^*}$ ; and  $\bar{u}_t = (1/N)\sum_{i=1}^N u_{i,t}$ .

Finally, the equation (4) is obtained through the equation (2) minus the equation (3).

$$(1/T)\Delta y_{i,t} = c_i - (1/T)(1 - e^{-\beta T})y_{i,t-T} + \varepsilon_{i,t} \quad (4)$$

where  $\Delta y_{i,t} = y_{i,t} - y_{i,t-T} = \log(Y_{i,t}/\bar{Y}_i) - \log(Y_{i,t-T}/\bar{Y}_{t-T})$ ;  $c_i = \alpha_i - \bar{\alpha} = (1/T)(1 - e^{-\beta T})y_i^*$  almost holds because both  $x_i$  and  $\bar{x}$  are positive and small enough so that the difference  $x_i - \bar{x}$  can be neglected,  $y_i^* = \log(Y_i^*/\bar{Y}^*)$ , so  $y_i^*$  denotes the relative steady state of per-capita output (log version) of economy  $i$  for all  $i$ ; and  $\varepsilon_{i,t} = u_{i,t} - \bar{u}_i$ .

In this paper, the equation (4) was used to test the hypothesis of  $\beta$  convergence. In the equation (4),  $c_i$  is the constant term of economy  $i$  for all  $i$ . In the case of conditional convergence,  $Y_i^*$  changes with  $i$ , then  $Y_i^*$  does not equal  $\bar{Y}^*$  for most  $i$  or  $y_i^*$  does not equal zero for most  $i$ , thus  $c_i$  does not equal zero for most  $i$  or  $c_i$  is significant for most  $i$ . While in the case of absolute convergence, the reverse occurs  $c_i$  is not significant for most  $i$ .

## 4. DATA, THE EMPIRICAL METHODOLOGY, THE RESULTS AND THE ANALYSES

### 4.1. The data

World Bank provides data on GDP per-capita (constant 2010 US\$) for countries and regions around the world. The downloaded data on GDP per-capita cover the years from 1970 to 2019 and includes 114 countries and regions<sup>4</sup> which are listed in *Appendix A* and whose data on GDP per-capita are available every year from 1970 to 2019.

### 4.2. The empirical methodology

Firstly, the above-mentioned data was regarded as a joint sample (the 1970-2019 sample), which consisted of the five sub-samples: the 1970-1979 sub-sample, the 1980-1989 sub-sample, the 1990-1999 sub-sample, the 2000-2009 sub-sample and the 2010-2019 sub-sample. There are both developed and less developed countries in each sub-sample, so conditional convergence should exist in each one.



Secondly, the regression results obtained by using the data in the five sub-samples could provide an estimate of relative steady state of per-capita output of each country in 1970s, 1980s, 1990s, 2000s and 2010s, respectively. If the hypothesis of conditional convergence was tested in the five sub-samples separately, the five estimates of each country would be obtained separately. According to econometrics, without making a Wald test, one cannot simply use two estimates to judge whether the change in a variable or the difference between two variables is significant. To make Wald tests for the assessments, the five estimates of all countries must be obtained simultaneously so that all estimates can be associated with each other in the econometric software. To solve this problem, dummy variables could be introduced into the regression equation.

In the equation (4),  $(1/T)(1 - e^{-\beta T}) \cong \beta$  holds when  $\beta$  is a very small positive number, so the constant term  $c_i = \beta y_i^*$  holds for country  $i$  for all  $i$ . Take one year as the time interval of observations used, i.e.,  $T = 1$  year, the equation (4) is rewritten as

$$\Delta y_{i,t} = c_i - \beta y_{i,t-1} + \varepsilon_{i,t} \tag{5}$$

Four dummy variables  $D1$ ,  $D2$ ,  $D3$  and  $D4$  were introduced into the equation (5) to capture, respectively, the changes in constant term  $c_i$  of country  $i$  for all  $i$  across sub-periods. Another four dummy variables  $DT1$ ,  $DT2$ ,  $DT3$  and  $DT4$  were introduced to find, respectively, the changes in the average speed of convergence  $\beta$  for all countries in the sample across sub-periods. In this way, the following equation can be obtained.

$$\begin{aligned} \Delta y_{i,t} = & c_{i,0} + \gamma_{i,1}D1 + \gamma_{i,2}D2 + \gamma_{i,3}D3 + \gamma_{i,4}D4 \\ & - \beta_0 y_{i,t-1} + \lambda_1 DT1 y_{i,t-1} + \lambda_2 DT2 y_{i,t-1} + \lambda_3 DT3 y_{i,t-1} + \lambda_4 DT4 y_{i,t-1} + \varepsilon_{i,t} \end{aligned} \tag{6}$$

where  $D1 = DT1 = 1$  when data is in the 1980-1989 sub-sample,  $D1 = DT1 = 0$  otherwise;  $D2 = DT2 = 1$  when data is in the 1990-1999 sub-sample,  $D2 = DT2 = 0$  otherwise;  $D3 = DT3 = 1$  when data is in the 2000-2009 sub-sample,  $D3 = DT3 = 0$  otherwise;  $D4 = DT4 = 1$  when data is in the 2010-2018 sub-sample,  $D4 = DT4 = 0$  otherwise;  $c_{i,0}$  denotes the constant term (fixed effect) of country  $i$  for all  $i$  in 1970s;  $\gamma_{i,1}$  denotes the gap between  $c_i$  in 1970s and 1980s for all  $i$ ;  $\gamma_{i,2}$  denotes the gap between  $c_i$  in 1970s and 1990s for all  $i$ ;  $\gamma_{i,3}$

denotes the gap between  $c_i$  in 1970s and 2000s for all  $i$ ;  $\gamma_{i,4}$  denotes the gap between  $c_i$  in 1970s and 2010s for all  $i$ ;  $\beta_0$  denotes the average speed of convergence for all countries in the sample in 1970s;  $\lambda_1$  denotes the gap between  $\beta$  in 1970s and 1980s;  $\lambda_2$  denotes the gap between  $\beta$  in 1970s and 1990s;  $\lambda_3$  denotes the gap between  $\beta$  in 1970s and 2000s; and  $\lambda_4$  denotes the gap between  $\beta$  in 1970s and 2010s. Further,  $c_{i,0}$ ,  $(c_{i,0} + \gamma_{i,1}) = c_{i,1}$ ,  $(c_{i,0} + \gamma_{i,2}) = c_{i,2}$ ,  $(c_{i,0} + \gamma_{i,3}) = c_{i,3}$  and  $(c_{i,0} + \gamma_{i,4}) = c_{i,4}$  denote the constant term of country  $i$  for all  $i$  in 1970s, 1980s, 1990s, 2000s, and 2010s, respectively;  $\beta_0$ ,  $(\beta_0 - \lambda_1) = \beta_1$ ,  $(\beta_0 - \lambda_2) = \beta_2$ ,  $(\beta_0 - \lambda_3) = \beta_3$ ,  $(\beta_0 - \lambda_4) = \beta_4$  denote the average speed of convergence for all countries in the sample in 1970s, 1980s, 1990s, 2000s and 2010s, respectively. After such an introduction of eight dummy variables, data in the five sub-samples was used jointly to estimate the equation (6) to obtain simultaneously the estimates of all above coefficients, this means it is feasible to obtain simultaneously the five estimates of relative steady states of per-capita output of all countries in the sample in the above five successive sub-periods.

### 4.3. The results and the analyses

According to the definition of conditional convergence, if  $\beta_0$  in the equation (6),  $\beta_1$ ,  $\beta_2$ ,  $\beta_3$  and  $\beta_4$ , which are implied in the equation (6), are all positive;  $c_{i,0}$  in the equation (6),  $c_{i,1}$ ,  $c_{i,2}$ ,  $c_{i,3}$  and  $c_{i,4}$ , which are implied in the equation (6), are all significant for most  $i$ , the hypothesis of conditional convergence cannot be rejected, respectively, in the 1970-1979 sub-sample, the 1980-1989 sub-sample, the 1990-1999 sub-sample, the 2000-2009 sub-sample and the 2010-2019 sub-sample.

Now make the following ten null hypotheses for the above five sub-samples:  $H_0: \beta_0 = 0$ ,  $H_0: c_{i,0} = 0$ ;  $H_0: \beta_1 = 0$ ,  $H_0: c_{i,1} = 0$ ;  $H_0: \beta_2 = 0$ ,  $H_0: c_{i,2} = 0$ ;  $H_0: \beta_3 = 0$ ,  $H_0: c_{i,3} = 0$ ;  $H_0: \beta_4 = 0$ ,  $H_0: c_{i,4} = 0$ . The regression results obtained from estimating the equation (6) using data in the five sub-samples jointly are shown in *Appendix B*, and the regression results about Brazil, Mexico, Malaysia, Turkey and South Africa, China and United States are selected and shown in Table 1.

**Table 1: The selected regression results obtained from estimating the equation (6)**

| Method: GLS (Cross Section Weights)         |                 |           |            |             |         |
|---|-----------------|-----------|------------|-------------|---------|
| Sample (adjusted): 1971 2019                |                 |           |            |             |         |
| Included observations: 49 after adjustments |                 |           |            |             |         |
| Number of cross-sections included: 114      |                 |           |            |             |         |
| Total pool (balanced) observations: 5586    |                 |           |            |             |         |
| Variable                                    | Coefficient     | Estimates | Std. Error | t-statistic | p value |
| $y_{i,t-1}$                                 | $-\beta_0$      | -0.193381 | 0.041258   | -4.687150   | 0.0000  |
| $DT1y_{i,t-1}$                              | $\lambda_1$     | 0.072969  | 0.049827   | 1.464443    | 0.1431  |
| $DT2y_{i,t-1}$                              | $\lambda_2$     | 0.000806  | 0.059595   | 0.013529    | 0.9892  |
| $DT3y_{i,t-1}$                              | $\lambda_3$     | 0.125194  | 0.047810   | 2.618578    | 0.0089  |
| $DT4y_{i,t-1}$                              | $\lambda_4$     | 0.044145  | 0.047281   | 0.933680    | 0.3505  |
| $c_0(BRA)$                                  | $c_0(BRA)$      | -0.038935 | 0.012127   | -3.210495   | 0.0013  |
| $D1(BRA)$                                   | $\gamma_1(BRA)$ | -0.004447 | 0.019268   | -0.230782   | 0.8175  |
| $D2(BRA)$                                   | $\gamma_2(BRA)$ | -0.057058 | 0.019029   | -2.998473   | 0.0027  |
| $D3(BRA)$                                   | $\gamma_3(BRA)$ | 0.010909  | 0.019224   | 0.567486    | 0.5704  |
| $D4(BRA)$                                   | $\gamma_4(BRA)$ | -0.036503 | 0.021156   | -1.725425   | 0.0845  |
| $c_0(CHN)$                                  | $c_0(CHN)$      | -0.673089 | 0.149284   | -4.508779   | 0.0000  |
| $D1(CHN)$                                   | $\gamma_1(CHN)$ | 0.358790  | 0.170019   | 2.110291    | 0.0349  |
| $D2(CHN)$                                   | $\gamma_2(CHN)$ | 0.258660  | 0.179484   | 1.441134    | 0.1496  |
| $D3(CHN)$                                   | $\gamma_3(CHN)$ | 0.587545  | 0.156993   | 3.973932    | 0.0001  |
| $D4(CHN)$                                   | $\gamma_4(CHN)$ | 0.562153  | 0.151754   | 3.704379    | 0.0002  |
| $c_0(MEX)$                                  | $c_0(MEX)$      | -0.063913 | 0.018300   | -3.492496   | 0.0005  |
| $D1(MEX)$                                   | $\gamma_1(MEX)$ | 0.012418  | 0.026273   | 0.472665    | 0.6365  |
| $D2(MEX)$                                   | $\gamma_2(MEX)$ | -0.021504 | 0.028940   | -0.743055   | 0.4575  |
| $D3(MEX)$                                   | $\gamma_3(MEX)$ | 0.013689  | 0.021505   | 0.636532    | 0.5245  |
| $D4(MEX)$                                   | $\gamma_4(MEX)$ | -0.024957 | 0.022820   | -1.093623   | 0.2742  |
| $c_0(MYS)$                                  | $c_0(MYS)$      | -0.223053 | 0.056818   | -3.925718   | 0.0001  |
| $D1(MYS)$                                   | $\gamma_1(MYS)$ | 0.109449  | 0.065681   | 1.666368    | 0.0957  |
| $D2(MYS)$                                   | $\gamma_2(MYS)$ | 0.092091  | 0.068271   | 1.348909    | 0.1774  |
| $D3(MYS)$                                   | $\gamma_3(MYS)$ | 0.185691  | 0.058642   | 3.166513    | 0.0016  |
| $D4(MYS)$                                   | $\gamma_4(MYS)$ | 0.165387  | 0.058394   | 2.832254    | 0.0046  |
| $c_0(TUR)$                                  | $c_0(TUR)$      | -0.123010 | 0.018541   | -6.634603   | 0.0000  |
| $D1(TUR)$                                   | $\gamma_1(TUR)$ | 0.048924  | 0.022437   | 2.180507    | 0.0293  |
| $D2(TUR)$                                   | $\gamma_2(TUR)$ | 0.020144  | 0.031600   | 0.637467    | 0.5238  |
| $D3(TUR)$                                   | $\gamma_3(TUR)$ | 0.094700  | 0.019641   | 4.821570    | 0.0000  |
| $D4(TUR)$                                   | $\gamma_4(TUR)$ | 0.100434  | 0.019974   | 5.028275    | 0.0000  |
| $c_0(USA)$                                  | $c_0(USA)$      | 0.199378  | 0.042692   | 4.670105    | 0.0000  |

|                     |                 |           |          |           |        |
|---------------------|-----------------|-----------|----------|-----------|--------|
| $D1(USA)$           | $\gamma_1(USA)$ | -0.064690 | 0.054067 | -1.196488 | 0.2316 |
| $D2(USA)$           | $\gamma_2(USA)$ | 0.014198  | 0.064549 | 0.219949  | 0.8259 |
| $D3(USA)$           | $\gamma_3(USA)$ | -0.129767 | 0.050865 | -2.551186 | 0.0108 |
| $D4(USA)$           | $\gamma_4(USA)$ | -0.040304 | 0.048932 | -0.823685 | 0.4102 |
| $c_0(ZAF)$          | $c_0(ZAF)$      | -0.088085 | 0.013689 | -6.434616 | 0.0000 |
| $D1(ZAF)$           | $\gamma_1(ZAF)$ | 0.009201  | 0.023762 | 0.387221  | 0.6986 |
| $D2(ZAF)$           | $\gamma_2(ZAF)$ | -0.087301 | 0.036817 | -2.371223 | 0.0178 |
| $D3(ZAF)$           | $\gamma_3(ZAF)$ | 0.035947  | 0.025307 | 1.420404  | 0.1556 |
| $D4(ZAF)$           | $\gamma_4(ZAF)$ | -0.053242 | 0.024067 | -2.212231 | 0.0270 |
| R-squared: 0.342232 |                 |           |          |           |        |

In Table 1, the p value of the t-statistic for the estimate of  $\beta_0$  shows  $H_0: \beta_0 = 0$  is rejected at the 1% significance level, and the estimate of  $\beta_0$  shows  $\beta_0$  is positive. In *Appendix B*, p values of t-statistics for most estimates of  $c_{i,0}$  show  $H_0: c_{i,0} = 0$  is rejected at the 1% significance level. The regression results of  $\beta_0$  and  $c_{i,0}$  show the hypothesis of conditional convergence is not rejected in the 1970-1979 sub-sample.

The regression results obtained from estimating the equation (6) do not provide directly the information about  $\beta_1, c_{i,1}, \beta_2, c_{i,2}, \beta_3, c_{i,3}, \beta_4$  and  $c_{i,4}$ , but Wald tests can be used to get the information about them. Table 2 contains the main results of all Wald tests made in this paper, and the original details are shown in *Appendix C*.

In Table 2, the results of the Wald test of  $H_0: \beta_1 = 0$  show the p value for the Chi-square is 0.0001, thus  $H_0: \beta_1 = 0$  is rejected at the 1% significance level, and the estimate of  $\beta_1$  ( $\hat{\beta}_1 = \hat{\beta}_0 - \hat{\lambda}_1 = 0.119674$ ) is calculated using the estimates of  $\beta_1$  and  $\lambda_1$  shown in Table 1, so  $\beta_1$  is positive. The Wald test of  $H_0: c_{i,1} = 0$  can be made on the country by country basis, but such a job is not done in this paper because of too many countries and districts in the sample. Since  $c_{i,0}$  is significant for most  $i$  while  $\gamma_{i,1}$  is not significant for most  $i$  according to p values of t-statistics for their estimates shown in *Appendix B*, one can infer  $c_{i,1} (= c_{i,0} + \gamma_{i,1})$  is significant for most  $i$ , that is, if the Wald test of  $H_0: c_{i,1} = 0$  is done, the results would show  $H_0: c_{i,1} = 0$  is rejected at the 5% or 10% significance level. Thus the information obtained about  $\beta_1$  and  $c_{i,1}$  show that the

hypothesis of conditional convergence is not rejected in the 1980-1989 sub-sample.

**Table 2: The results of all Wald tests made in this paper**

|  |          |         |        |
|--|----------|---------|--------|
| 1. Null Hypothesis: $\beta_1 = 0$                  |          |         |        |
| Chi-square   | 18.57513 | p value | 0.0000 |
| 2. Null Hypothesis: $\beta_2 = 0$                  |          |         |        |
| Chi-square   | 20.05227 | p value | 0.0000 |
| 3. Null Hypothesis: $\beta_3 = 0$                  |          |         |        |
| Chi-square   | 7.966911 | p value | 0.0048 |
| 4. Null Hypothesis: $\beta_4 = 0$                  |          |         |        |
| Chi-square   | 41.76206 | p value | 0.0000 |
| 5. Null Hypothesis: $y_1^*(BRA) - y_0^*(BRA) = 0$  |          |         |        |
| Chi-square   | 2.168572 | p value | 0.1409 |
| 6. Null Hypothesis: $y_2^*(BRA) - y_1^*(BRA) = 0$  |          |         |        |
| Chi-square   | 1.367203 | p value | 0.2423 |
| 7. Null Hypothesis: $y_3^*(BRA) - y_2^*(BRA) = 0$  |          |         |        |
| Chi-square   | 0.569852 | p value | 0.4503 |
| 8. Null Hypothesis: $y_4^*(BRA) - y_3^*(BRA) = 0$  |          |         |        |
| Chi-square   | 0.587786 | p value | 0.4433 |
| 9. Null Hypothesis: $y_0^*(BRA) = 0$               |          |         |        |
| Chi-square   | 65.09192 | p value | 0.0000 |
| 10. Null Hypothesis: $y_1^*(BRA) = 0$              |          |         |        |
| Chi-square   | 11.77179 | p value | 0.0006 |
| 11. Null Hypothesis: $y_2^*(BRA) = 0$              |          |         |        |
| Chi-square   | 84.46333 | p value | 0.0000 |
| 12. Null Hypothesis: $y_3^*(BRA) = 0$              |          |         |        |
| Chi-square   | 16.11826 | p value | 0.0001 |
| 13. Null Hypothesis: $y_4^*(BRA) = 0$              |          |         |        |
| Chi-square   | 54.30023 | p value | 0.0000 |
| 14. Null Hypothesis: $y_0^*(CHN) - y_0^*(BRA) = 0$ |          |         |        |
| Chi-square   | 1357.977 | p value | 0.0000 |
| 15. Null Hypothesis: $y_1^*(CHN) - y_1^*(BRA) = 0$ |          |         |        |
| Chi-square   | 229.4352 | p value | 0.0000 |
| 16. Null Hypothesis: $y_2^*(CHN) - y_2^*(BRA) = 0$ |          |         |        |
| Chi-square   | 251.7001 | p value | 0.0000 |
| 17. Null Hypothesis: $y_3^*(CHN) - y_3^*(BRA) = 0$ |          |         |        |
| Chi-square   | 126.0329 | p value | 0.0000 |
| 18. Null Hypothesis: $y_4^*(CHN) - y_4^*(BRA) = 0$ |          |         |        |
| Chi-square   | 5.523785 | p value | 0.0188 |

Similarly, using the above method, one can know that  $\beta_2, \beta_3$  and  $\beta_4$  are all positive;  $H_0 : c_{i,2} = 0, H_0 : c_{i,3} = 0$  and  $H_0 : c_{i,4} = 0$  are all rejected at the 5% or

10% significance level. So the information obtained about  $\beta_2$ ,  $c_{i,2}$ ,  $\beta_3$ ,  $c_{i,3}$ ,  $\beta_4$  and  $c_{i,4}$  suggest the hypothesis of conditional convergence is not rejected in the 1990-1999 sub-sample, the 2000-2009 sub-sample and the 2010-2019 sub-sample.

As shown in Section 3,  $y_i^* = \log(Y_i^* / Y^*)$  denotes the relative steady state of per-capita output (log version) of country  $i$  for all  $i$ . Let  $y_{i,0}^*$ ,  $y_{i,1}^*$ ,  $y_{i,2}^*$ ,  $y_{i,3}^*$  and  $y_{i,4}^*$  denote the relative steady state of per-capita output of country  $i$  for all  $i$  in 1970s, 1980s, 1990s, 2000s and 2010s, respectively. As shown in Section 4.2,  $c_i = \beta y_i^*$  holds for country  $i$  for all  $i$ , so the estimate of  $y_i^*$  can be computed by using the estimates of  $c_i$  and  $\beta$  in each sub-period. Now take Brazil's relative steady state of per-capita output  $y^*(BRA)$  as an example. Using the concerned estimates provided in Table 1, the details of the computation are shown as follows.

$$\hat{y}_0^*(BRA) = \hat{c}_0(BRA) / \hat{\beta}_0 = -0.038935 / 0.193381 = -0.2013$$

$$\begin{aligned} \hat{y}_1^*(BRA) &= \hat{c}_1(BRA) / \hat{\beta}_1 = [\hat{c}_0(BRA) + \hat{\gamma}_1(BRA)] / (\hat{\beta}_0 - \hat{\lambda}_1) \\ &= (-0.038935 - 0.004447) / (0.193381 - 0.072969) = -0.3605 \end{aligned}$$

$$\begin{aligned} \hat{y}_2^*(BRA) &= \hat{c}_2(BRA) / \hat{\beta}_2 = [\hat{c}_0(BRA) + \hat{\gamma}_2(BRA)] / (\hat{\beta}_0 - \hat{\lambda}_2) \\ &= (-0.038935 - 0.057058) / (0.193381 - 0.000806) = -0.4984 \end{aligned}$$

$$\begin{aligned} \hat{y}_3^*(BRA) &= \hat{c}_3(BRA) / \hat{\beta}_3 = [\hat{c}_0(BRA) + \hat{\gamma}_3(BRA)] / (\hat{\beta}_0 - \hat{\lambda}_3) \\ &= (-0.038935 + 0.010909) / (0.193381 - 0.125194) = -0.4106 \end{aligned}$$

$$\begin{aligned} \hat{y}_4^*(BRA) &= \hat{c}_4(BRA) / \hat{\beta}_4 = [\hat{c}_0(BRA) + \hat{\gamma}_4(BRA)] / (\hat{\beta}_0 - \hat{\lambda}_4) \\ &= (-0.038935 - 0.036503) / (0.193381 - 0.044145) = -0.5054 \end{aligned}$$

Similarly, let  $y^*(MEX)$ ,  $y^*(MYS)$ ,  $y^*(TUR)$ ,  $y^*(ZAF)$ ,  $y^*(CHN)$  and  $y^*(USA)$  denote, respectively, relative steady states of per-capita output of Mexico, Malaysia, Turkey, South Africa, China and United States, one can compute their estimates by using the above method. All estimates of the seven countries are shown in Table 3.

**Table 3: The estimates of relative steady states of per-capita output of the five “middle-income trap” countries, China and United States**

| Names of countries | Estimates in 1970s | Estimates in 1980s | Estimates in 1990s | Estimates in 2000s | Estimates in 2010s |
|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|
| Brazil             | -0.2013            | -0.3605            | -0.4984            | -0.4106            | -0.5054            |
| Mexico             | -0.3305            | -0.4277            | -0.4434            | -0.7364            | -0.5956            |
| Malaysia           | -1.1534            | -0.9435            | -0.6802            | -0.5484            | -0.3867            |
| Turkey             | -0.6361            | -0.6153            | -0.5343            | -0.4150            | -0.1515            |
| South Africa       | -0.4555            | -0.6553            | -0.9107            | -0.7639            | -0.9471            |
| China              | -3.4806            | -2.6105            | -2.1516            | -1.2543            | -0.7433            |
| United States      | 1.0310             | 1.1188             | 1.1090             | 1.0205             | 1.0662             |

In Table 3, the estimates of United States are all positive. United States is a typical developed country; its steady state of per-capita output  $Y(USA)$  is always higher than the average  $\bar{Y}^*$  of all sample countries, so its relative steady state of per-capita output  $y^*(USA)$  is always significantly positive, actually around 1. The estimates of Brazil, Mexico, Malaysia, Turkey, South Africa and China are all significantly negative or near to zero as shown in Table 3 because they are all less developed countries.

How to assess whether a country’s relative steady state of per-capita output  $y^*$  changes with time? Take  $y^*(BRA)$  as an example and make the four null hypotheses:  $H_0: y_1^*(BRA) - y_0^*(BRA) = 0$ ,  $H_0: y_2^*(BRA) - y_1^*(BRA) = 0$ ,  $H_0: y_3^*(BRA) - y_2^*(BRA) = 0$ ,  $H_0: y_4^*(BRA) - y_3^*(BRA) = 0$ . In Table 2, the results of the Wald test of  $H_0: y_1^*(BRA) - y_0^*(BRA) = 0$  show that the p value for the Chi-square is above 10%, which means  $H_0: y_1^*(BRA) - y_0^*(BRA) = 0$  is not rejected, i.e., the gap between  $y^*(BRA)$  in 1970s and 1980s is possibly not significant. Thus Brazil’s relative steady state of per-capita output possibly did not change significantly from 1970s to 1980s.

Similarly, according to the results of the Wald tests of  $H_0: y_2^*(BRA) - y_1^*(BRA) = 0$ ,  $H_0: y_3^*(BRA) - y_2^*(BRA) = 0$  and  $H_0: y_4^*(BRA) - y_3^*(BRA) = 0$ , all of the three null hypotheses are not rejected because their p values for the Chi-square are all above 10% as shown in Table 2. So it is possible that Brazil’s relative steady state of per-capita output did not change significantly from 1980s to 1990s, from 1990s to 2000s, and from 2000s to 2010s.

The formula  $y_i^* = \log(Y_i^* / \bar{Y}^*)$  shows 0 is the average of relative steady states of per-capita output (log version) of all countries in the test sample. Now make

five null hypotheses:  $H_0: y_0^*(BRA) = 0$ ,  $H_0: y_1^*(BRA) = 0$ ,  $H_0: y_2^*(BRA) = 0$ ,  $H_0: y_3^*(BRA) = 0$ ,  $H_0: y_4^*(BRA) = 0$ . In Table 2, the results of the Wald tests show all above five null hypotheses are rejected at the 1% significance level because their p values for the Chi-square are all below 1%. As shown in Table 3, the five estimates of  $y^*(BRA)$  are all negative, so Brazil's relative steady state of per-capita output is significantly below the average of all countries in the test sample in each of the five sub-periods.

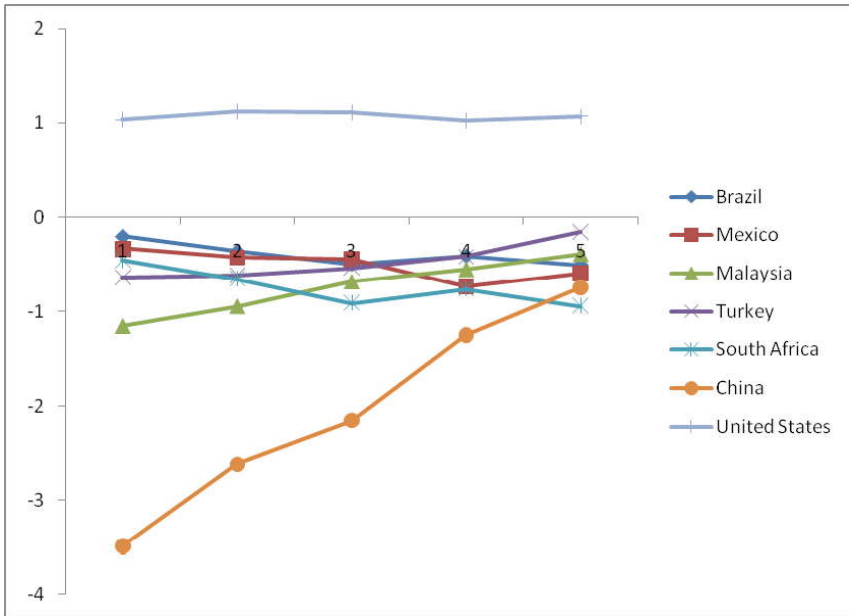
How to assess whether a country's relative steady state of per-capita output differs from another's in the same period? Take  $y^*(CHN)$  and  $y^*(BRA)$  as an example and make five null hypotheses:  $H_0: y_0^*(CHN) - y_0^*(BRA) = 0$ ,  $H_0: y_1^*(CHN) - y_1^*(BRA) = 0$ ,  $H_0: y_2^*(CHN) - y_2^*(BRA) = 0$ ,  $H_0: y_3^*(CHN) - y_3^*(BRA) = 0$  and  $H_0: y_4^*(CHN) - y_4^*(BRA) = 0$ . In Table 2, the results of Wald tests show all above null hypothesis are rejected at the 1% or 5% significance level according to their p values for the Chi-square. The estimates of  $y^*(BRA)$  and  $y^*(CHN)$  are shown in Table 3, so one can judge China's relative steady state of per-capita output is lower than Brazil's in each sub-period.

## 5. THE PATHS OF RELATIVE STEADY STATES OF PER-CAPITA OUTPUT OF THE SEVEN COUNTRIES

The path of relative steady state of per-capita output of a country shows how the steady state of per-capita output of the country changes relatively in a test sample, i.e., in terms of steady state of per-capita output, it shows how the relative position of a country changes in a test sample. The path is obtained by using the estimates of the relative steady state of per-capita output of a country in some successive sub-periods. In this paper, the paths of Brazil, Mexico, Malaysia, Turkey, South Africa, China and United States are drawn by using their estimates in Table 3 and shown in Figure 2.

As shown earlier,  $y_i^* = \log(Y_i^* / Y^*)$  is the formula for the relative steady state of per-capita output (log version) of country  $i$  for all  $i$ . In Figure 2, the horizontal axis is for such a hypothetical country: its relative steady state of per-capita output always equals 0, i.e., its steady state of per-capita output always equals the average level of all countries in a test sample. A description of Figure 2 is shown as follows.





**Figure 2: The paths of relative steady states of per-capita output of the five “middle-income trap” countries, China and United States (1970-2019)**

*Note:* 1. The numbers 1, 2, 3, 4 and 5 below the horizontal axis denote 1970s, 1980s, 1990s, 2000s and 2010s, respectively. 2. The numbers at the left side of the vertical axis denote the measures of relative steady state of per-capita output.

In Figure 2, the path of the US is obviously above the horizontal axis, so it is a typical path of a developed country. Due to capital accumulation and technological progress, it is reasonable to believe the US’s steady state of per-capita output kept growing from 1970s to 2010s, but the path of the US shows the US’s relative steady state of per-capita output did not change significantly from 1970s to 2010s, i.e., in terms of steady state of per-capita output, the US’s relative position in the test sample did not change significantly in the 1970-2019 period.

The paths of Brazil, Mexico, Malaysia, Turkey and South Africa are all below the horizontal axis, but generally not far apart. The five paths show that the relative steady states of per-capita output of the five countries generally fluctuated slightly from 1970s to 2010s, i.e., in terms of steady state of per-capita output, their relative positions in the test sample did not change greatly on the whole from 1970s to 2010s. The above situation shows that even in terms of steady state of per-capita output, the five countries were not only developing countries but also typical “middle income trap” countries in the 1970-2019 period.

The path of China is generally far below the horizontal axis. The path shows China's relative steady state of per-capita output was extremely low in 1970s, then it kept increasing dramatically, and almost caught up with the overall level of the above five countries in 2010s, i.e., in terms of steady state of per-capita output, China's relative position in the test sample kept rising significantly after 1970s, and almost reached the overall level of the above five countries in 2010s. The path of China suggests, in terms of steady state of per-capita output, China was a developing country but not a "middle income trap" country in the 1970-2019 period, or to be more exact, China started to face the "middle income trap" in 2010s.

## **6. AN ANALYSIS OF THE REASONS FOR THE RELATIVE CHANGES IN STEADY STATES OF PER-CAPITA OUTPUT OF THE SEVEN COUNTRIES**

The Solow model shows that an economy's steady state of per-capita output depends on its economic parameters and effectiveness of labour. To be precise, it is an economy's social infrastructure that determines its steady state of per-capita output through influencing the economic parameters and the effectiveness of labour. As Romer described<sup>5</sup>, the social infrastructure refers to those institutions, policies, traditions and cultures, which can influence economic growth. Next, by looking up historical data of the five typical "middle income trap" countries, this paper will reveal how their social infrastructures determined their steady states of per-capita output ( $Af(k^*)$ ) through influencing the saving rate ( $s$ ), the population growth rate ( $n$ ) and the effectiveness of labour ( $A$ ) in the 1970-2019 period. In addition, in view of the need of the research made in this paper, the data of China, the United States and the world are also looked up.

First, look at the saving rate. The data on annual saving rates of the concerned countries and the world were downloaded from the World Bank database, and their average annual saving rates in the 1970s, 1980s, 1990s, 2000s and 2010 were calculated, respectively, and shown in Table 4. The data in Table 4 provides the eight paths in Figure 3, which reflects roughly the changes in the saving rates of the seven countries and the world in the 1970-2019 period.

Table 4 and Figure 3 show that in the 1970-2019 period, the world's saving rate did not change significantly. Among the five typical "middle income trap" countries, only Malaysia's saving rate was higher than the world level in each sub-period while the saving rates of other four countries were lower

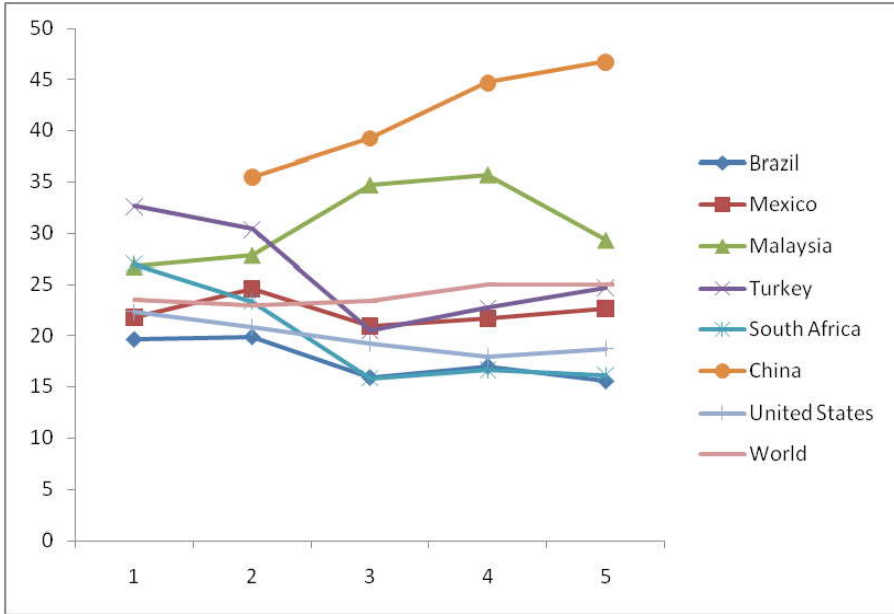
than the world's at least in most sub-periods, thus, except Malaysia, the saving rates of the other four countries were generally lower than the world level, i.e., their saving rates might be lower than the average level of all sample countries. In addition, in the 1970-2019 period, Malaysia's saving situation was obviously better than that of other four countries, but Malaysia's saving rate was not high enough because the country did not keep its saving rate above 30% in this period. The convergence theory shows, other factors remain unchanged, a higher saving rate leads to a higher  $k^*$  and  $f(k^*)$ , and the converse is also true. The above situations show that the saving rates of the five countries were generally not high in the 1970-2019 period, this should be an important reason to explain why in this period, in terms of the steady state of per-capita output, the relative positions of the five countries in the test sample were always slightly lower than the average level of the test sample and also did not change greatly as a whole.

**Table 4: The saving rates of the concerned countries and the world in the 1970-2019 period (%)**

| <i>Names of countries</i> | <i>Averages in 1970s</i> | <i>Averages in 1980s</i> | <i>Averages in 1990s</i> | <i>Averages in 2000s</i> | <i>Averages in 2010s</i> |
|---------------------------|--------------------------|--------------------------|--------------------------|--------------------------|--------------------------|
| Brazil                    | 19.61                    | 19.86                    | 15.87                    | 16.97                    | 15.51                    |
| Mexico                    | 21.82                    | 24.61                    | 20.93                    | 21.7                     | 22.68                    |
| Malaysia                  | 26.75                    | 27.81                    | 34.73                    | 35.72                    | 29.29                    |
| Turkey                    | 32.62                    | 30.33                    | 20.48                    | 22.71                    | 24.7                     |
| South Africa              | 27                       | 23.32                    | 15.78                    | 16.59                    | 16.07                    |
| China                     | —                        | 35.43                    | 39.21                    | 44.66                    | 46.74                    |
| United States             | 22.32                    | 20.83                    | 19.17                    | 17.89                    | 18.7                     |
| World                     | 23.46                    | 22.98                    | 23.43                    | 24.98                    | 24.96                    |

*Note:* The World Bank database lacks data on China's annual saving rate in 1970s, so China's average annual saving rate in 1970s is blank in Table 4.

Second, look at the population growth rate. The data on annual population growth rates of the concerned countries and the world are downloaded from the World Bank database, and their average annual population growth rates in the 1970s, 1980s, 1990s, 2000s and 2010 are calculated, respectively, and listed in Table 5. The data in Table 5 generate the eight paths in Figure 4, which reflects basically the changes in the population growth rates of the seven countries and the world in the 1970-2019 period.

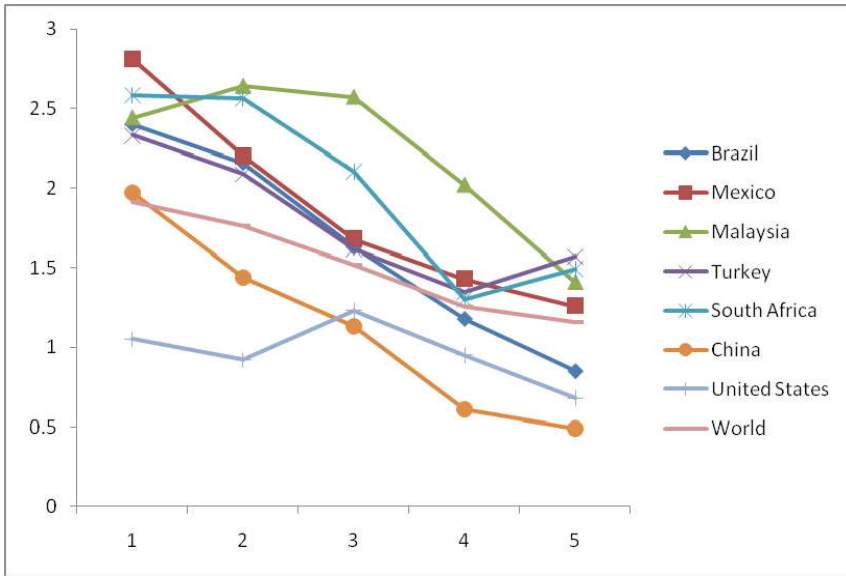


**Figure 3: The paths of the saving rates of the concerned countries and the world (1970-2019)**

*Note:* 1. The numbers 1, 2, 3, 4 and 5 below the horizontal axis denote 1970s, 1980s, 1990s, 2000s and 2010s, respectively. 2. The numbers at the left side of the vertical axis denote the measures (%) of saving rate. 3. The World Bank database lacks data on China’s annual saving rate in 1970s, so there is one corresponding blank for China’s path in Figure 3.

**Table 5: The population growth rates of the concerned countries and the world in the 1970-2019 period (%)**

| <i>Names of countries</i> | <i>Averages in 1970s</i> | <i>Averages in 1980s</i> | <i>Averages in 1990s</i> | <i>Averages in 2000s</i> | <i>Averages in 2010s</i> |
|---------------------------|--------------------------|--------------------------|--------------------------|--------------------------|--------------------------|
| Brazil                    | 2.4                      | 2.16                     | 1.63                     | 1.18                     | 0.85                     |
| Mexico                    | 2.81                     | 2.2                      | 1.68                     | 1.43                     | 1.26                     |
| Malaysia                  | 2.44                     | 2.64                     | 2.57                     | 2.02                     | 1.41                     |
| Turkey                    | 2.33                     | 2.09                     | 1.62                     | 1.35                     | 1.57                     |
| South Africa              | 2.58                     | 2.56                     | 2.1                      | 1.3                      | 1.49                     |
| China                     | 1.97                     | 1.44                     | 1.13                     | 0.61                     | 0.49                     |
| United States             | 1.05                     | 0.92                     | 1.23                     | 0.95                     | 0.68                     |
| World                     | 1.91                     | 1.76                     | 1.52                     | 1.26                     | 1.16                     |



**Figure 4: The paths of the population growth rates of the concerned countries and the world (1970-2019)**

*Note:* 1. The numbers 1, 2, 3, 4 and 5 below the horizontal axis denote 1970s, 1980s, 1990s, 2000s and 2010s, respectively. 2. The numbers at the left side of the vertical axis denote the measures (%) of population growth rate.

Table 5 and Figure 4 show that in the 1970-2019 period, the world’s population growth rate kept an obvious downward trend, and the population growth rates of the five typical “middle income trap” countries also showed a similar downward trend as a whole. According to the theory of convergence, other factors remain unchanged, a lower population growth rate leads to a higher level of  $k^*$  and  $f(k^*)$ , and the converse is also true. A downward trend in population growth rates of the five countries should be helpful to increase their  $f(k^*)$ . However, the world’s population growth rate maintained a marked downward trend in the 1970-2019 period, it means most sample countries also experienced a similar downward trend in their population growth rates in this period, so logically the population growth rates of the five countries might not show a relative decline in the test sample. Thus, although the population growth rates of the five countries had a general downward trend, this might make no significant effect on improving the relative positions of the steady states of per-capita output of the five countries in the test sample. On the other hand, as shown in Table 5 and Figure 4, during the 1970-2019 period, the population

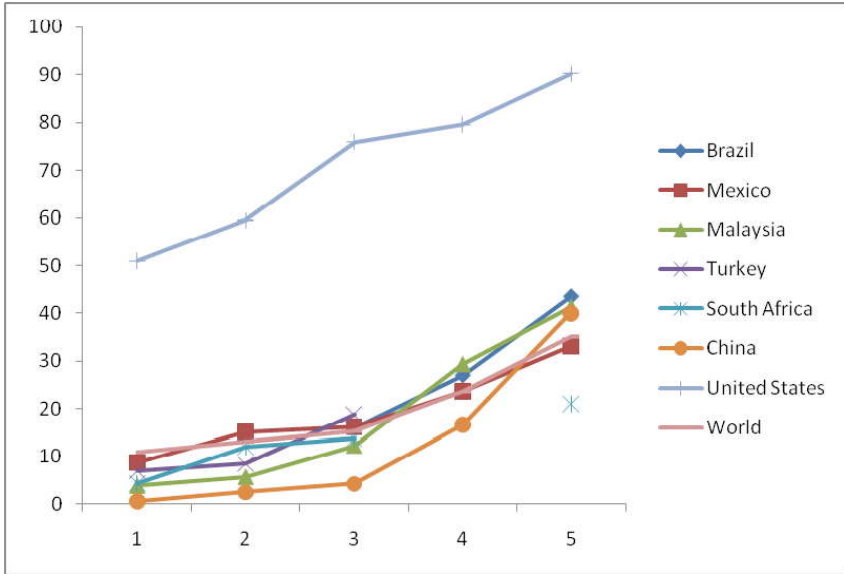
growth rates of the five countries were actually higher than the world level (except Brazil in the 2000s and 2010s), it means that the population growth rates of the five countries might be higher than the average level of the test sample in this period. This situation is helpful to explain why in the 1970-2019 period, their relative positions of the steady states of per-capita output were always slightly lower than the average level of the test sample.

Finally, look at the labour efficiency (i.e., the effectiveness of labour mentioned in the Section 2). Labour efficiency ( $A$ ) undoubtedly makes a huge effect on the steady state of per-capita output  $Af(k^*)$ . Human capital is the source of technological progress and innovation, so one can think of human capital as the most important indicator to measure labour efficiency. Unfortunately, since 2017 the World Bank's database has only begun to provide data on the human capital index of countries and regions around the world, thus this paper chose the tertiary school enrollment rate to roughly reflect the human capital level. The data on annual tertiary school enrollment rates of the concerned countries and the world were downloaded from the World Bank database, and their average annual tertiary school enrollment rate in 1970s, 1980s, 1990s, 2000s and 2010 were calculated, respectively, and listed in Table 6. The data in Table 6 give the eight paths in Figure 5, which shows the changes in the tertiary school enrollment rates of the seven countries and the world in the 1970-2019 period, and also roughly reflects the changes in the human capital and the labour efficiency of the seven countries and the world in this period.

**Table 6: The tertiary school enrollment rates of the concerned countries and the world in the 1970-2019 period (%)**

| <i>Names of countries</i> | <i>Averages in 1970s</i> | <i>Averages in 1980s</i> | <i>Averages in 1990s</i> | <i>Averages in 2000s</i> | <i>Averages in 2010s</i> |
|---------------------------|--------------------------|--------------------------|--------------------------|--------------------------|--------------------------|
| Brazil                    | —                        | —                        | 16.08                    | 26.99                    | 43.46                    |
| Mexico                    | 8.78                     | 15.13                    | 16.26                    | 23.75                    | 33.2                     |
| Malaysia                  | 3.82                     | 5.69                     | 12.32                    | 29.4                     | 41.36                    |
| Turkey                    | 6.87                     | 8.34                     | 18.9                     | —                        | —                        |
| South Africa              | 4.33                     | 12                       | 13.82                    | —                        | 20.99                    |
| China                     | 0.5                      | 2.44                     | 4.21                     | 16.61                    | 40.09                    |
| United States             | 50.9                     | 59.45                    | 75.9                     | 79.52                    | 90.14                    |
| World                     | 10.95                    | 13.1                     | 15.39                    | 23.82                    | 35.22                    |

*Note:* The World Bank database lacks the data on Brazil's annual tertiary school enrollment rate in 1970s and 1980s, the data on Turkey's in 2000s and 2010s, and the data on South Africa 's in 2000s, so there are some corresponding blanks for them in Table 6.



**Figure 5: The paths of the tertiary school enrollment rates of the concerned countries and the world (1970-2019)**

*Note:* 1. The numbers 1, 2, 3, 4 and 5 below the horizontal axis denote 1970s, 1980s, 1990s, 2000s and 2010s, respectively. 2. The numbers at the left side of the vertical axis denote the measures (%) of tertiary school enrollment rate. 3. The World Bank database lacks the data on Brazil in 1970s and 1980s, the data on Turkey’s in 2000s and 2010s, and the data on South Africa ‘s in 2000s, so there are some corresponding blanks for them in Figure 5.

Table 6 and Figure 5 show that in the 1970-2019 period, the tertiary school enrolment rate of the world kept an obvious upward trend, so did the tertiary school enrolment rates of the five typical “middle income trap” countries, and both had a less difference as a whole. It can be inferred that in the 1970-2019 period, the human capital and labour efficiency of the five countries were constantly improving, which should be very helpful to increase the steady states of per-capita output of the five countries. However, the tertiary school enrolment rate of the world kept an obvious upward trend, it means that in this period, the tertiary school enrolment rates of most sample countries had a similar upward trend, accordingly, the human capital of most sample countries had a similar upward trend, so the human capital of the five countries might show no relative improvement in the test sample, and so did the labour efficiency of the five countries. Thus, the tertiary school enrolment rates of the five countries kept rising in the 1970-2019 period, but it might play no significant role in improving

the relative positions of steady states of per-capita output of the five countries in the test sample.

China's steady state of per-capita output maintained a significant relative growth in the test sample because the changes in China's indicators were completely different in the 1970-2019 period. As shown in Table 4 and Figure 3, in the 1970-2019 period, except in 1970s (the data on China's saving rate in 1970s is not available), China's saving rate kept rising and was much higher than the overall level of the five countries and the world level, and should be also much higher than average level of all sample countries. Table 5 and Figure 4 show that, China's population growth rate experienced a downward trend like the five countries' and the world's in the 1970-2019 period, but was lower than the overall level of the five countries, also lower than the world level (except almost equal the world level in the 1970s), and should be also lower than the average level of all sample countries. Finally, there happened a dramatic growth in China's tertiary school enrollment rate in the 1970-2019 period. As shown in Table 6 and Figure 5, China's tertiary school enrollment rate was extremely low (0.5%) in 1970s, but by 2010s it was higher than the world level and even caught up with the high-end level of the five countries, so it grew much faster than the five countries' and the world's, undoubtedly also much faster than most sample countries', and so did China's human capital and labour efficiency in this period. The above changes in China's indicators can explain why China's steady state of per-capita output kept growing rapidly and relatively in the test sample in the 1970-2019 period.

Finally, a brief analysis of the reasons for the situation in the United States is given. As shown in Table 4 and Figure 3, in the 1970-2019 period, the US's saving rate showed a slight downward trend, and was always lower than the world level, and might be also lower than the average level of all sample countries, so this certainly made a negative effect on the relative position of the US's steady state of per-capita output in the test sample. However, Table 5 and Figure 4 show that during this period, the US's population growth rate also had a slight downward trend and was always lower than the world level, and might be also lower than the average level of all sample countries. In addition, as shown in Table 6 and Figure 5, during this period, the US's tertiary school enrollment rate kept an obvious upward trend like the world's, more importantly, it was much higher than the world level, and should be also much higher than the average level of all sample countries, and so was US's human capital and labour efficiency in this period. The above changes in the US's population growth rate and tertiary school enrollment rate (especially the latter) surely made a



significantly greater positive effect on the relative position of the US's steady state of per-capita output in the test sample, and actually played a dominant role. Thus, the net result from the changes in the US's indicators is that during the 1970-2019 period, the relative position of the US's steady state of per-capita output in the test sample did not change significantly, and were always significantly higher than the average level of the test sample, and certainly also significantly higher than the relative positions of the five typical "middle income trap" countries and China.

## **7. CONCLUSIONS**

Based on the theory of convergence, this paper used econometric method mainly revealing the followings: (1) In terms of steady state of per-capita output, the relative positions of Brazil, Mexico, Malaysia, Turkey and South Africa in the test sample generally remained slightly below the average level of the sample countries in the 1970-2019 period, i.e., in terms of steady state of per-capita output, the five countries were typical "middle income trap" countries in this period. (2) China's relative position in the test sample was far below the overall level of the five countries in 1970s, but kept rising rapidly since then, and almost reached the overall level of the five countries in 2010s, so in terms of steady state of per-capita output, China was not a "middle income trap" country in the 1970-2019 period, but started to face the "middle income trap" in 2010s.

This paper also provides an analysis of the reasons for the above situations. In general, the social infrastructures of the five typical "middle-income trap" countries did not change significantly in the 1970-2019 period, this resulted in the followings: in this period, their saving rates (except Malaysia's) should be lower than the average level of the sample countries; their population growth rates should be higher than the average level of the sample countries; their human capital did not significantly exceed the most sample countries', and nor did their labour efficiency. Therefore, if the five countries want to make their steady states of per-capita output increase relatively in the test sample in the future, their governments must formulate feasible and effective policies to improve their social infrastructures so as to significantly increase their saving rates, reduce further their population growth rates, and achieve a quicker growth of their human capital and labour efficiency, or the five countries will continue to stay in the "middle-income trap".

China's social infrastructure was improved significantly in the 1970-2019 period due to its many correct policies executed since the late 1970s, this resulted in the followings: in this period, China's saving rate maintained at a high level

and kept rising except in 1970s, actually much higher than the average level of the sample countries except in 1970s; China's population growth rate kept declining, actually lower than the average level of the sample countries; China's tertiary school enrollment rate was extremely low in 1970s, but it grew much faster than most sample countries', and so did China's human capital and labour efficiency. But it is necessary to point out, China's saving rate was already very high (46.74% in 2010s) and much difficult to increase further significantly, its population growth rate was already very low (0.49% in 2010s) and leaves little room to decrease further, but China's tertiary school enrollment rate was not high (40.09% in 2010s) in comparison with developed countries (e.g. the United States). So the Chinese government should pay more attention to promoting the growth of China's human capital and labour efficiency in the future. The future growth of China's human capital and labour efficiency will, to large extent, determine the future growth of China's steady state of per-capita output, which will decide whether China can smoothly cross the "middle income trap" after 2010s.

### *Notes*

1. The typical "middle income trap" countries selected in this paper refer to those that were always "middle income trap" countries in the 1970-2019 period. After meeting this requirement and considering the geographical distribution of the selected countries, Brazil and Mexico were selected as representative countries in the America; Malaysia and Turkey were selected as representative countries in Asia; South Africa was chosen as the representative country in Africa. Since the World Bank database had not provided the per-capita GDP data of Eastern European countries in 1970s and 1980s, in order to ensure the rigour of the regression results, the test sample used in this paper does not include the "middle income trap" countries in Europe (such as Bulgaria and Romania, etc.).
2. For more details of the Solow model, see Romer (2001, Chapter 1).
3. Romer, D. 2001. *Advanced Macroeconomics*. 2nd edition. New York: McGraw-Hill. P.21
4. World Bank provides data on GDP per-capita of countries and regions in the world from 1960 to 2019, but data in 1960s are not available for most countries, even data in 1970s are not available for some countries, so this paper has to choose the 114 countries and regions and a data time span from 1970 to 2019 to form a test sample.
5. See Romer (2001, p.143)

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## Appendix A

### The 114 Countries and Regions with Their Codes in the Sample

#### 1. The name list of 29 developed countries and regions

Andorra—AND, Australia—AUS, Austria—AUT, Belgium—BEL, Canada—CAN, Switzerland—CHE, Germany—DEU, Denmark—DNK, Spain—ESP, Finland—FIN, France—FRA, United Kingdom—GBR, Greece—GRC, Greenland—GRL, Hong Kong—HKG, Ireland—IRL, Iceland—ISL, Israel—ISR, Italy—ITA, Japan—JPN, Luxembourg—LUX, Monaco—MCO, Netherlands—NLD, Norway—NOR, New Zealand—NZL, Portugal—PRT, Singapore—SGP, Sweden—SWE, United States—USA

#### 2. The name list of 85 developing countries

Argentina—ARG, Burundi—BDI, Benin—BEN, Burkina Faso—BFA, Bangladesh—BGD, Bahamas—BHS, Belize—BLZ, Bolivia—BOL, Brazil—BRA, Botswana—BWA, Central African Republic—CAF, Chile—CHL, China—CHN, Cote d'Ivoire—CIV, Cameroon—CMR, Congo, Dem. Rep.—COD, Congo, Republic of—COG, Colombia—COL, Costa Rica—CRI, Cuba—CUB, Dominican Republic—DOM, Algeria—DZA, Ecuador—ECU, Egypt—EGY, Fiji—FJI, Gabon—GAB, Georgia—GEO, Ghana—GHA, Gambia—GMB, Guinea-Bissau—GNB, Guatemala—GTM, Guyana—GUY, Honduras—HND, Haiti—HTI, Indonesia—IDN, India—IND, Iran—IRN, Iraq—IRQ, Jamaica—JAM, Kenya—KEN, Kiribati—KIR, Korea, Republic of—KOR, Sri Lanka—LKA, Lesotho—LSO, Morocco—MAR, Madagascar—MDG, Mexico—MEX, Mali—MLI, Malta—MLT, Myanmar—MMR, Mauritania—MRT, Malawi—MWI, Malaysia—MYS, Niger—NER, Nigeria—NGA, Nicaragua—NIC, Nepal—NPL, Oman—OMN, Pakistan—PAK, Panama—PAN, Peru—PER, Philippines—PHL, Papua New Guinea—PNG, Puerto Rico—PRI, Paraguay—PRY, Rwanda—RWA, Saudi Arabia—SAU, Sudan—SDN, Senegal—SEN, Sierra Leone—SLE, El Salvador—SLV, Suriname—SUR, Swaziland—SWZ, Seychelles—SYC, Chad—TCD, Togo—TGO, Thailand—THA, Trinidad & Tobago—TTO, Tunisia—TUN, Turkey—TUR, Uruguay—URY, St. Vincent and the Grenadines—VCT, South Africa—ZAF, Zambia—ZMB, Zimbabwe—ZWE

## Appendix B

### The Regression Results from Estimating the Equation (6) (Outputs of Eviews)

Dependent Variable: D(Y?)

Method: Pooled EGLS (Cross-section weights)

Date: 08/01/21 Time: 23:20

Sample (adjusted): 1971 2019

Included observations: 49 after adjustments

Cross-sections included: 114

Total pool (balanced) observations: 5586

Linear estimation after one-step weighting matrix

White cross-section standard errors & covariance (d.f. corrected)

| <i>Variable</i> | <i>Coefficient</i> | <i>Std. Error</i> | <i>t-Statistic</i> | <i>Prob.</i> |
|-----------------|--------------------|-------------------|--------------------|--------------|
| 1 Y?(-1)        | C(1)= -0.193381    | 0.041258          | -4.687150          | 0.0000       |
| DT1*Y?(-1)      | C(2)= 0.072969     | 0.049827          | 1.464443           | 0.1431       |
| DT2*Y?(-1)      | C(3)= 0.000806     | 0.059595          | 0.013529           | 0.9892       |
| DT3*Y?(-1)      | C(4)= 0.125194     | 0.047810          | 2.618578           | 0.0089       |
| 5 DT4*Y?(-1)    | C(5)= 0.044145     | 0.047281          | 0.933680           | 0.3505       |
| AND—C           | 0.271399           | 0.069710          | 3.893285           | 0.0001       |
| ARG—C           | -0.050771          | 0.016708          | -3.038728          | 0.0024       |
| AUS—C           | 0.201802           | 0.047824          | 4.219635           | 0.0000       |
| AUT—C           | 0.190616           | 0.038714          | 4.923712           | 0.0000       |
| 10 BDI—C        | -0.699684          | 0.142295          | -4.917156          | 0.0000       |
| BEL—C           | 0.184124           | 0.039325          | 4.682124           | 0.0000       |
| BEN—C           | -0.501747          | 0.102466          | -4.896714          | 0.0000       |
| BFA—C           | -0.653104          | 0.134012          | -4.873475          | 0.0000       |
| BGD—C           | -0.665966          | 0.135407          | -4.918261          | 0.0000       |
| 15 BHS—C        | 0.140782           | 0.020893          | 6.738232           | 0.0000       |
| BLZ—C           | -0.311833          | 0.071099          | -4.385875          | 0.0000       |
| BOL—C           | -0.339140          | 0.063980          | -5.300713          | 0.0000       |
| BRA—C           | C(18)= -0.038935   | 0.012127          | -3.210495          | 0.0013       |
| BWA—C           | -0.297091          | 0.085622          | -3.469783          | 0.0005       |
| 20 CAF—C        | -0.544847          | 0.103422          | -5.268208          | 0.0000       |
| CAN—C           | 0.201966           | 0.044869          | 4.501262           | 0.0000       |
| CHE—C           | 0.318145           | 0.068980          | 4.612162           | 0.0000       |
| CHL—C           | -0.156884          | 0.048447          | -3.238220          | 0.0012       |
| CHN—C           | C(24)= -0.673089   | 0.149284          | -4.508779          | 0.0000       |
| 25 CIV—C        | -0.280008          | 0.057898          | -4.836193          | 0.0000       |
| CMR—C           | -0.398326          | 0.093412          | -4.264170          | 0.0000       |
| COD—C           | -0.470284          | 0.088137          | -5.335855          | 0.0000       |
| COG—C           | -0.304220          | 0.064920          | -4.686077          | 0.0000       |
| COL—C           | -0.187820          | 0.042145          | -4.456568          | 0.0000       |
| 30 CRI—C        | -0.131716          | 0.029746          | -4.428057          | 0.0000       |
| CUB—C           | -0.223799          | 0.049619          | -4.510358          | 0.0000       |
| DEU—C           | 0.177455           | 0.036377          | 4.878222           | 0.0000       |
| DNK—C           | 0.244714           | 0.052524          | 4.659069           | 0.0000       |
| DOM—C           | -0.261452          | 0.055407          | -4.718707          | 0.0000       |
| 35 DZA—C        | -0.202663          | 0.046211          | -4.385593          | 0.0000       |
| ECU—C           | -0.192870          | 0.037151          | -5.191515          | 0.0000       |
| EGY—C           | -0.453020          | 0.097865          | -4.629028          | 0.0000       |
| ESP—C           | 0.110332           | 0.028792          | 3.832081           | 0.0001       |

|    |       |                  |          |           |        |
|----|-------|------------------|----------|-----------|--------|
|    | FIN—C | 0.168876         | 0.035662 | 4.735498  | 0.0000 |
| 40 | FJI—C | -0.233823        | 0.056344 | -4.149930 | 0.0000 |
|    | FRA—C | 0.186346         | 0.037971 | 4.907618  | 0.0000 |
|    | GAB—C | 0.078445         | 0.064909 | 1.208535  | 0.2269 |
|    | GBR—C | 0.150140         | 0.030877 | 4.862539  | 0.0000 |
|    | GEO—C | -0.189259        | 0.046061 | -4.108848 | 0.0000 |
| 45 | GHA—C | -0.472373        | 0.098635 | -4.789097 | 0.0000 |
|    | GMB—C | -0.482186        | 0.088236 | -5.464720 | 0.0000 |
|    | GNB—C | -0.562635        | 0.115898 | -4.854572 | 0.0000 |
|    | GRC—C | 0.127991         | 0.028601 | 4.474998  | 0.0000 |
|    | GRL—C | 0.179885         | 0.030051 | 5.986066  | 0.0000 |
| 50 | GTM—C | -0.262862        | 0.056500 | -4.652457 | 0.0000 |
|    | GUY—C | -0.312539        | 0.052282 | -5.977967 | 0.0000 |
|    | HKG—C | -0.006477        | 0.018846 | -0.343699 | 0.7311 |
|    | HND—C | -0.360510        | 0.077647 | -4.642942 | 0.0000 |
|    | HTI—C | -0.423921        | 0.094045 | -4.507648 | 0.0000 |
| 55 | IDN—C | -0.423826        | 0.094092 | -4.504394 | 0.0000 |
|    | IND—C | -0.625306        | 0.118810 | -5.263059 | 0.0000 |
|    | IRL—C | 0.099270         | 0.023988 | 4.138409  | 0.0000 |
|    | IRN—C | -0.034777        | 0.036355 | -0.956588 | 0.3388 |
|    | IRQ—C | -0.271517        | 0.077482 | -3.504237 | 0.0005 |
| 60 | ISL—C | 0.183444         | 0.030789 | 5.958185  | 0.0000 |
|    | ISR—C | 0.117838         | 0.026272 | 4.485221  | 0.0000 |
|    | ITA—C | 0.160096         | 0.032233 | 4.966903  | 0.0000 |
|    | JAM—C | -0.173314        | 0.035177 | -4.926955 | 0.0000 |
|    | JPN—C | 0.175763         | 0.035879 | 4.898706  | 0.0000 |
| 65 | KEN—C | -0.459318        | 0.109019 | -4.213192 | 0.0000 |
|    | KIR—C | -0.222343        | 0.056360 | -3.945033 | 0.0001 |
|    | KOR—C | -0.191524        | 0.055517 | -3.449855 | 0.0006 |
|    | LKA—C | -0.484168        | 0.101000 | -4.793754 | 0.0000 |
|    | LSO—C | -0.561408        | 0.138618 | -4.050025 | 0.0001 |
| 70 | LUX—C | 0.277305         | 0.058583 | 4.733566  | 0.0000 |
|    | MAR—C | -0.394786        | 0.080972 | -4.875576 | 0.0000 |
|    | MCO—C | 0.457498         | 0.097240 | 4.704822  | 0.0000 |
|    | MDG—C | -0.518442        | 0.104345 | -4.968532 | 0.0000 |
|    | MEX—C | C(74)= -0.063913 | 0.018300 | -3.492496 | 0.0005 |
| 75 | MLI—C | -0.601184        | 0.127735 | -4.706498 | 0.0000 |
|    | MLT—C | -0.042741        | 0.032967 | -1.296480 | 0.1949 |
|    | MMR—C | -0.773176        | 0.163889 | -4.717689 | 0.0000 |

|           |                   |          |           |        |
|-----------|-------------------|----------|-----------|--------|
| MRT—C     | -0.346585         | 0.067239 | -5.154557 | 0.0000 |
| MWI—C     | -0.614457         | 0.135777 | -4.525493 | 0.0000 |
| 80 MYS—C  | C(80)= -0.223053  | 0.056818 | -3.925718 | 0.0001 |
| NER—C     | -0.534001         | 0.120557 | -4.429467 | 0.0000 |
| NGA—C     | -0.300006         | 0.066508 | -4.510849 | 0.0000 |
| NIC—C     | -0.321612         | 0.039191 | -8.206229 | 0.0000 |
| NLD—C     | 0.210572          | 0.047298 | 4.452011  | 0.0000 |
| 85 NOR—C  | 0.293799          | 0.061562 | 4.772422  | 0.0000 |
| NPL—C     | -0.692232         | 0.139466 | -4.963453 | 0.0000 |
| NZL—C     | 0.153572          | 0.038108 | 4.029935  | 0.0001 |
| OMN—C     | -0.012654         | 0.037946 | -0.333471 | 0.7388 |
| PAK—C     | -0.583550         | 0.118343 | -4.931001 | 0.0000 |
| 90 PAN—C  | -0.179718         | 0.040711 | -4.414481 | 0.0000 |
| PER—C     | -0.196846         | 0.032592 | -6.039677 | 0.0000 |
| PHL—C     | -0.346043         | 0.072419 | -4.778335 | 0.0000 |
| PNG—C     | -0.345865         | 0.071360 | -4.846790 | 0.0000 |
| PRI—C     | 0.066429          | 0.007922 | 8.385786  | 0.0000 |
| 95 PRT—C  | 0.037652          | 0.009959 | 3.780528  | 0.0002 |
| PRY—C     | -0.252540         | 0.064933 | -3.889237 | 0.0001 |
| RWA—C     | -0.635407         | 0.141297 | -4.496966 | 0.0000 |
| SAU—C     | 0.278707          | 0.057613 | 4.837575  | 0.0000 |
| SDN—C     | -0.487128         | 0.081195 | -5.999491 | 0.0000 |
| 100 SEN—C | -0.420861         | 0.078811 | -5.340101 | 0.0000 |
| SGP—C     | 0.044131          | 0.005716 | 7.721103  | 0.0000 |
| SLE—C     | -0.595720         | 0.120473 | -4.944850 | 0.0000 |
| SLV—C     | -0.235943         | 0.041661 | -5.663361 | 0.0000 |
| SUR—C     | -0.066407         | 0.021288 | -3.119462 | 0.0018 |
| 105 SWE—C | 0.212940          | 0.050946 | 4.179702  | 0.0000 |
| SWZ—C     | -0.353707         | 0.068763 | -5.143854 | 0.0000 |
| SYC—C     | -0.095222         | 0.049677 | -1.916831 | 0.0553 |
| TCD—C     | -0.593813         | 0.089860 | -6.608216 | 0.0000 |
| TGO—C     | -0.529967         | 0.102767 | -5.156974 | 0.0000 |
| 110 THA—C | -0.395992         | 0.088591 | -4.469899 | 0.0000 |
| TTO—C     | -0.064014         | 0.019766 | -3.238636 | 0.0012 |
| TUN—C     | -0.313436         | 0.071780 | -4.366632 | 0.0000 |
| TUR—C     | C(113)= -0.123010 | 0.018541 | -6.634603 | 0.0000 |
| URY—C     | -0.083577         | 0.019583 | -4.267810 | 0.0000 |
| 115 USA—C | C(115)= 0.199378  | 0.042692 | 4.670105  | 0.0000 |
| VCT—C     | -0.296802         | 0.080495 | -3.687212 | 0.0002 |



|            |                   |          |           |        |
|------------|-------------------|----------|-----------|--------|
| ZAF—C      | -0.088085         | 0.013689 | -6.434616 | 0.0000 |
| ZMB—C      | -0.391928         | 0.067164 | -5.835347 | 0.0000 |
| ZWE—C      | -0.408702         | 0.078503 | -5.206173 | 0.0000 |
| 120 AND—D1 | -0.150468         | 0.077860 | -1.932541 | 0.0533 |
| ARG—D1     | -0.027725         | 0.020727 | -1.337639 | 0.1811 |
| AUS—D1     | -0.068247         | 0.056716 | -1.203308 | 0.2289 |
| AUT—D1     | -0.065700         | 0.048976 | -1.341480 | 0.1798 |
| BDI—D1     | 0.270036          | 0.173786 | 1.553846  | 0.1203 |
| 125 BEL—D1 | -0.062075         | 0.047411 | -1.309302 | 0.1905 |
| BEN—D1     | 0.182536          | 0.124766 | 1.463024  | 0.1435 |
| BFA—D1     | 0.246131          | 0.161827 | 1.520952  | 0.1283 |
| BGD—D1     | 0.259215          | 0.162390 | 1.596249  | 0.1105 |
| BHS—D1     | -0.025573         | 0.041160 | -0.621314 | 0.5344 |
| 130 BLZ—D1 | 0.129148          | 0.091558 | 1.410567  | 0.1584 |
| BOL—D1     | 0.059123          | 0.087103 | 0.678771  | 0.4973 |
| BRA—D1     | C(132)= -0.004447 | 0.019268 | -0.230782 | 0.8175 |
| BWA—D1     | 0.185933          | 0.095463 | 1.947697  | 0.0515 |
| CAF—D1     | 0.154345          | 0.130510 | 1.182631  | 0.2370 |
| 135 CAN—D1 | -0.072806         | 0.054879 | -1.326651 | 0.1847 |
| CHE—D1     | -0.114570         | 0.082543 | -1.388003 | 0.1652 |
| CHL—D1     | 0.067926          | 0.060518 | 1.122402  | 0.2617 |
| CHN—D1     | C(138)= 0.358790  | 0.170019 | 2.110291  | 0.0349 |
| CIV—D1     | 0.010277          | 0.074906 | 0.137196  | 0.8909 |
| 140 CMR—D1 | 0.160804          | 0.103428 | 1.554743  | 0.1201 |
| COD—D1     | 0.139599          | 0.110356 | 1.264985  | 0.2059 |
| COG—D1     | 0.165365          | 0.078109 | 2.117110  | 0.0343 |
| COL—D1     | 0.067595          | 0.050064 | 1.350182  | 0.1770 |
| CRI—D1     | 0.011633          | 0.038932 | 0.298804  | 0.7651 |
| 145 CUB—D1 | 0.121336          | 0.056940 | 2.130959  | 0.0331 |
| DEU—D1     | -0.059420         | 0.045822 | -1.296742 | 0.1948 |
| DNK—D1     | -0.082304         | 0.068600 | -1.199772 | 0.2303 |
| DOM—D1     | 0.096246          | 0.068894 | 1.397026  | 0.1625 |
| DZA—D1     | 0.060387          | 0.052339 | 1.153778  | 0.2486 |
| 150 ECU—D1 | 0.051007          | 0.048279 | 1.056506  | 0.2908 |
| EGY—D1     | 0.221880          | 0.113058 | 1.962529  | 0.0498 |
| ESP—D1     | -0.036955         | 0.032684 | -1.130700 | 0.2582 |
| FIN—D1     | -0.037151         | 0.044433 | -0.836121 | 0.4031 |
| FJI—D1     | 0.048018          | 0.072487 | 0.662438  | 0.5077 |
| 155 FRA—D1 | -0.066267         | 0.046946 | -1.411570 | 0.1581 |

|            |                  |          |           |        |
|------------|------------------|----------|-----------|--------|
| GAB—D1     | -0.093060        | 0.069893 | -1.331474 | 0.1831 |
| GBR—D1     | -0.043260        | 0.041364 | -1.045832 | 0.2957 |
| GEO—D1     | 0.078207         | 0.052330 | 1.494495  | 0.1351 |
| GHA—D1     | 0.133860         | 0.124724 | 1.073247  | 0.2832 |
| 160 GMB—D1 | 0.163593         | 0.112073 | 1.459704  | 0.1444 |
| GNB—D1     | 0.199735         | 0.142005 | 1.406535  | 0.1596 |
| GRC—D1     | -0.073037        | 0.032711 | -2.232766 | 0.0256 |
| GRL—D1     | -0.064291        | 0.037281 | -1.724502 | 0.0847 |
| GTM—D1     | 0.048273         | 0.073442 | 0.657288  | 0.5110 |
| 165 GUY—D1 | 0.054938         | 0.076969 | 0.713767  | 0.4754 |
| HKG—D1     | 0.071228         | 0.022838 | 3.118758  | 0.0018 |
| HND—D1     | 0.107720         | 0.093119 | 1.156791  | 0.2474 |
| HTI—D1     | 0.109538         | 0.113733 | 0.963117  | 0.3355 |
| IDN—D1     | 0.195014         | 0.112014 | 1.740975  | 0.0817 |
| 170 IND—D1 | 0.266688         | 0.147950 | 1.802558  | 0.0715 |
| IRL—D1     | -0.019751        | 0.027972 | -0.706090 | 0.4802 |
| IRN—D1     | -0.137593        | 0.052357 | -2.627955 | 0.0086 |
| IRQ—D1     | 0.095926         | 0.093365 | 1.027431  | 0.3043 |
| ISL—D1     | -0.059218        | 0.043569 | -1.359154 | 0.1742 |
| 175 ISR—D1 | -0.048499        | 0.033540 | -1.446035 | 0.1482 |
| ITA—D1     | -0.042954        | 0.040820 | -1.052299 | 0.2927 |
| JAM—D1     | 0.036232         | 0.047775 | 0.758395  | 0.4483 |
| JPN—D1     | -0.032809        | 0.046208 | -0.710016 | 0.4777 |
| KEN—D1     | 0.146700         | 0.130533 | 1.123850  | 0.2611 |
| 180 KIR—D1 | -0.030573        | 0.067945 | -0.449969 | 0.6528 |
| KOR—D1     | 0.156079         | 0.059309 | 2.631625  | 0.0085 |
| LKA—D1     | 0.212319         | 0.118740 | 1.788097  | 0.0738 |
| LSO—D1     | 0.198896         | 0.160009 | 1.243033  | 0.2139 |
| LUX—D1     | -0.072746        | 0.072470 | -1.003800 | 0.3155 |
| 185 MAR—D1 | 0.163127         | 0.097125 | 1.679568  | 0.0931 |
| MCO—D1     | -0.178938        | 0.116757 | -1.532576 | 0.1254 |
| MDG—D1     | 0.132547         | 0.130797 | 1.013377  | 0.3109 |
| MEX—D1     | C(188)= 0.012418 | 0.026273 | 0.472665  | 0.6365 |
| MLI—D1     | 0.209894         | 0.157226 | 1.334985  | 0.1819 |
| 190 MLT—D1 | 0.041920         | 0.034528 | 1.214089  | 0.2248 |
| MMR—D1     | 0.287226         | 0.194705 | 1.475183  | 0.1402 |
| MRT—D1     | 0.102148         | 0.086308 | 1.183520  | 0.2367 |
| MWI—D1     | 0.173107         | 0.160478 | 1.078695  | 0.2808 |
| MYS—D1     | C(194)= 0.109449 | 0.065681 | 1.666368  | 0.0957 |

|     |        |                   |          |           |        |
|-----|--------|-------------------|----------|-----------|--------|
| 195 | NER—D1 | 0.141961          | 0.148678 | 0.954822  | 0.3397 |
|     | NGA—D1 | 0.014175          | 0.091281 | 0.155291  | 0.8766 |
|     | NIC—D1 | 0.039145          | 0.065891 | 0.594089  | 0.5525 |
|     | NLD—D1 | -0.083304         | 0.056429 | -1.476250 | 0.1399 |
|     | NOR—D1 | -0.090963         | 0.079098 | -1.150005 | 0.2502 |
| 200 | NPL—D1 | 0.267154          | 0.170641 | 1.565594  | 0.1175 |
|     | NZL—D1 | -0.056273         | 0.046973 | -1.197991 | 0.2310 |
|     | OMN—D1 | 0.061183          | 0.043088 | 1.419963  | 0.1557 |
|     | PAK—D1 | 0.259839          | 0.142426 | 1.824388  | 0.0682 |
|     | PAN—D1 | 0.056655          | 0.052949 | 1.069999  | 0.2847 |
| 205 | PER—D1 | 0.025931          | 0.048371 | 0.536087  | 0.5919 |
|     | PHL—D1 | 0.096451          | 0.094087 | 1.025127  | 0.3054 |
|     | PNG—D1 | 0.085444          | 0.086278 | 0.990335  | 0.3221 |
|     | PRI—D1 | -0.016419         | 0.012352 | -1.329226 | 0.1838 |
|     | PRT—D1 | -0.000182         | 0.012372 | -0.014713 | 0.9883 |
| 210 | PRY—D1 | 0.109832          | 0.078196 | 1.404572  | 0.1602 |
|     | RWA—D1 | 0.226638          | 0.167982 | 1.349185  | 0.1773 |
|     | SAU—D1 | -0.283980         | 0.064026 | -4.435352 | 0.0000 |
|     | SDN—D1 | 0.159652          | 0.112483 | 1.419343  | 0.1559 |
|     | SEN—D1 | 0.129840          | 0.099346 | 1.306954  | 0.1913 |
| 215 | SGP—D1 | 0.042824          | 0.013783 | 3.107003  | 0.0019 |
|     | SLE—D1 | 0.189970          | 0.147402 | 1.288789  | 0.1975 |
|     | SLV—D1 | -0.003212         | 0.056274 | -0.057074 | 0.9545 |
|     | SUR—D1 | -0.017829         | 0.030109 | -0.592172 | 0.5538 |
|     | SWE—D1 | -0.069524         | 0.061795 | -1.125074 | 0.2606 |
| 220 | SWZ—D1 | 0.174157          | 0.090756 | 1.918943  | 0.0550 |
|     | SYC—D1 | 0.019358          | 0.054815 | 0.353157  | 0.7240 |
|     | TCD—D1 | 0.226127          | 0.121777 | 1.856895  | 0.0634 |
|     | TGO—D1 | 0.165135          | 0.135637 | 1.217479  | 0.2235 |
|     | THA—D1 | 0.205764          | 0.105518 | 1.950042  | 0.0512 |
| 225 | TTO—D1 | -0.024297         | 0.033923 | -0.716224 | 0.4739 |
|     | TUN—D1 | 0.110380          | 0.085021 | 1.298265  | 0.1943 |
|     | TUR—D1 | C(227)= 0.048924  | 0.022437 | 2.180507  | 0.0293 |
|     | URY—D1 | 0.009931          | 0.033152 | 0.299568  | 0.7645 |
|     | USA—D1 | C(229)= -0.064690 | 0.054067 | -1.196488 | 0.2316 |
| 230 | VCT—D1 | 0.154434          | 0.089315 | 1.729105  | 0.0839 |
|     | ZAF—D1 | C(231)= 0.009201  | 0.023762 | 0.387221  | 0.6986 |
|     | ZMB—D1 | 0.094684          | 0.093194 | 1.015985  | 0.3097 |
|     | ZWE—D1 | 0.154888          | 0.103903 | 1.490698  | 0.1361 |

|     |        |                   |          |           |        |
|-----|--------|-------------------|----------|-----------|--------|
|     | AND—D2 | -0.083429         | 0.080807 | -1.032448 | 0.3019 |
| 235 | ARG—D2 | -0.041565         | 0.023335 | -1.781212 | 0.0749 |
|     | AUS—D2 | 0.010245          | 0.066759 | 0.153464  | 0.8780 |
|     | AUT—D2 | 0.014908          | 0.058819 | 0.253465  | 0.7999 |
|     | BDI—D2 | -0.078691         | 0.218492 | -0.360155 | 0.7187 |
|     | BEL—D2 | 0.008960          | 0.057502 | 0.155815  | 0.8762 |
| 240 | BEN—D2 | -0.018589         | 0.156489 | -0.118789 | 0.9054 |
|     | BFA—D2 | 0.000204          | 0.198690 | 0.001027  | 0.9992 |
|     | BGD—D2 | 0.028456          | 0.196651 | 0.144702  | 0.8850 |
|     | BHS—D2 | -0.005438         | 0.038771 | -0.140268 | 0.8885 |
|     | BLZ—D2 | 0.064187          | 0.093828 | 0.684093  | 0.4939 |
| 245 | BOL—D2 | -0.075736         | 0.110656 | -0.684424 | 0.4937 |
|     | BRA—D2 | C(246)= -0.057058 | 0.019029 | -2.998473 | 0.0027 |
|     | BWA—D2 | 0.104852          | 0.099104 | 1.057999  | 0.2901 |
|     | CAF—D2 | -0.129269         | 0.177673 | -0.727565 | 0.4669 |
|     | CAN—D2 | -0.016750         | 0.063051 | -0.265654 | 0.7905 |
| 250 | CHE—D2 | -0.021882         | 0.096912 | -0.225789 | 0.8214 |
|     | CHL—D2 | 0.079042          | 0.050979 | 1.550476  | 0.1211 |
|     | CHN—D2 | C(252)= 0.258660  | 0.179484 | 1.441134  | 0.1496 |
|     | CIV—D2 | -0.168611         | 0.108915 | -1.548098 | 0.1217 |
|     | CMR—D2 | -0.103992         | 0.139146 | -0.747359 | 0.4549 |
| 255 | COD—D2 | -0.266082         | 0.169081 | -1.573694 | 0.1156 |
|     | COG—D2 | -0.034766         | 0.092005 | -0.377866 | 0.7055 |
|     | COL—D2 | -0.006936         | 0.055844 | -0.124194 | 0.9012 |
|     | CRI—D2 | -0.028754         | 0.044407 | -0.647516 | 0.5173 |
|     | CUB—D2 | -0.080367         | 0.085059 | -0.944833 | 0.3448 |
| 260 | DEU—D2 | 0.014048          | 0.055838 | 0.251582  | 0.8014 |
|     | DNK—D2 | 0.014351          | 0.078508 | 0.182800  | 0.8550 |
|     | DOM—D2 | 0.004801          | 0.076836 | 0.062478  | 0.9502 |
|     | DZA—D2 | -0.072752         | 0.073153 | -0.994529 | 0.3200 |
|     | ECU—D2 | -0.054254         | 0.060089 | -0.902881 | 0.3666 |
| 265 | EGY—D2 | 0.064870          | 0.131843 | 0.492025  | 0.6227 |
|     | ESP—D2 | 0.014698          | 0.038300 | 0.383752  | 0.7012 |
|     | FIN—D2 | 0.009753          | 0.053618 | 0.181898  | 0.8557 |
|     | FJI—D2 | -0.039054         | 0.086064 | -0.453779 | 0.6500 |
|     | FRA—D2 | 0.000219          | 0.056531 | 0.003873  | 0.9969 |
| 270 | GAB—D2 | -0.119177         | 0.066680 | -1.787303 | 0.0739 |
|     | GBR—D2 | 0.019032          | 0.049491 | 0.384565  | 0.7006 |
|     | GEO—D2 | -0.312313         | 0.132808 | -2.351606 | 0.0187 |

|     |        |                   |          |           |        |
|-----|--------|-------------------|----------|-----------|--------|
|     | GHA—D2 | -0.046954         | 0.150703 | -0.311565 | 0.7554 |
|     | GMB—D2 | -0.067330         | 0.148767 | -0.452589 | 0.6509 |
| 275 | GNB—D2 | -0.039309         | 0.182283 | -0.215648 | 0.8293 |
|     | GRC—D2 | -0.041990         | 0.035431 | -1.185122 | 0.2360 |
|     | GRL—D2 | -0.051187         | 0.051369 | -0.996453 | 0.3191 |
|     | GTM—D2 | -0.065584         | 0.090212 | -0.726996 | 0.4673 |
|     | GUY—D2 | -0.025070         | 0.092181 | -0.271959 | 0.7857 |
| 280 | HKG—D2 | 0.103076          | 0.031192 | 3.304609  | 0.0010 |
|     | HND—D2 | -0.061787         | 0.115402 | -0.535410 | 0.5924 |
|     | HTI—D2 | -0.143453         | 0.156450 | -0.916928 | 0.3592 |
|     | IDN—D2 | 0.076408          | 0.124263 | 0.614890  | 0.5387 |
|     | IND—D2 | 0.070686          | 0.174900 | 0.404154  | 0.6861 |
| 285 | IRL—D2 | 0.096763          | 0.040768 | 2.373517  | 0.0177 |
|     | IRN—D2 | -0.163333         | 0.061730 | -2.645915 | 0.0082 |
|     | IRQ—D2 | -0.005200         | 0.162795 | -0.031939 | 0.9745 |
|     | ISL—D2 | -0.010530         | 0.049487 | -0.212783 | 0.8315 |
|     | ISR—D2 | 0.010569          | 0.039864 | 0.265131  | 0.7909 |
| 290 | ITA—D2 | 0.017420          | 0.052145 | 0.334067  | 0.7383 |
|     | JAM—D2 | -0.024596         | 0.055705 | -0.441539 | 0.6588 |
|     | JPN—D2 | 0.039769          | 0.061141 | 0.650439  | 0.5154 |
|     | KEN—D2 | -0.082251         | 0.159293 | -0.516350 | 0.6056 |
|     | KIR—D2 | -0.190724         | 0.099599 | -1.914922 | 0.0556 |
| 295 | KOR—D2 | 0.202775          | 0.057820 | 3.507022  | 0.0005 |
|     | LKA—D2 | 0.082293          | 0.138037 | 0.596167  | 0.5511 |
|     | LSO—D2 | 0.015549          | 0.183730 | 0.084632  | 0.9326 |
|     | LUX—D2 | 0.074947          | 0.093690 | 0.799944  | 0.4238 |
|     | MAR—D2 | 0.017005          | 0.117055 | 0.145272  | 0.8845 |
| 300 | MCO—D2 | -0.021487         | 0.138426 | -0.155222 | 0.8767 |
|     | MDG—D2 | -0.134654         | 0.175142 | -0.768823 | 0.4420 |
|     | MEX—D2 | C(302)= -0.021504 | 0.028940 | -0.743055 | 0.4575 |
|     | MLI—D2 | -0.020670         | 0.187801 | -0.110066 | 0.9124 |
|     | MLT—D2 | 0.083618          | 0.033426 | 2.501597  | 0.0124 |
| 305 | MMR—D2 | 0.026738          | 0.237291 | 0.112681  | 0.9103 |
|     | MRT—D2 | -0.074729         | 0.109494 | -0.682496 | 0.4950 |
|     | MWI—D2 | -0.070966         | 0.206632 | -0.343441 | 0.7313 |
|     | MYS—D2 | C(308)= 0.092091  | 0.068271 | 1.348909  | 0.1774 |
|     | NER—D2 | -0.131076         | 0.181438 | -0.722429 | 0.4701 |
| 310 | NGA—D2 | -0.140789         | 0.118697 | -1.186119 | 0.2356 |
|     | NIC—D2 | -0.152782         | 0.112509 | -1.357949 | 0.1745 |

|            |                   |          |           |        |
|------------|-------------------|----------|-----------|--------|
| NLD—D2     | 0.010817          | 0.066299 | 0.163161  | 0.8704 |
| NOR—D2     | 0.041413          | 0.097136 | 0.426342  | 0.6699 |
| NPL—D2     | 0.026745          | 0.204150 | 0.131008  | 0.8958 |
| 315 NZL—D2 | -0.021313         | 0.050164 | -0.424855 | 0.6710 |
| OMN—D2     | 0.051046          | 0.041124 | 1.241265  | 0.2146 |
| PAK—D2     | 0.040851          | 0.166809 | 0.244898  | 0.8065 |
| PAN—D2     | 0.004888          | 0.058465 | 0.083610  | 0.9334 |
| PER—D2     | -0.091712         | 0.066455 | -1.380048 | 0.1676 |
| 320 PHL—D2 | -0.068956         | 0.115879 | -0.595068 | 0.5518 |
| PNG—D2     | -0.039093         | 0.105500 | -0.370554 | 0.7110 |
| PRI—D2     | 0.025483          | 0.025382 | 1.003978  | 0.3154 |
| PRT—D2     | 0.038263          | 0.017621 | 2.171452  | 0.0299 |
| PRY—D2     | 0.006001          | 0.081862 | 0.073302  | 0.9416 |
| 325 RWA—D2 | -0.084146         | 0.220187 | -0.382156 | 0.7024 |
| SAU—D2     | -0.209871         | 0.062853 | -3.339075 | 0.0008 |
| SDN—D2     | -0.034637         | 0.136347 | -0.254037 | 0.7995 |
| SEN—D2     | -0.075993         | 0.132813 | -0.572180 | 0.5672 |
| SGP—D2     | 0.121293          | 0.033879 | 3.580221  | 0.0003 |
| 330 SLE—D2 | -0.145929         | 0.201829 | -0.723032 | 0.4697 |
| SLV—D2     | -0.079125         | 0.080503 | -0.982883 | 0.3257 |
| SUR—D2     | -0.112173         | 0.037571 | -2.985647 | 0.0028 |
| SWE—D2     | -0.005879         | 0.069054 | -0.085143 | 0.9322 |
| SWZ—D2     | 0.075419          | 0.099125 | 0.760850  | 0.4468 |
| 335 SYC—D2 | 0.026874          | 0.054018 | 0.497502  | 0.6189 |
| TCD—D2     | -0.055569         | 0.161049 | -0.345046 | 0.7301 |
| TGO—D2     | -0.094360         | 0.173483 | -0.543916 | 0.5865 |
| THA—D2     | 0.142940          | 0.108880 | 1.312814  | 0.1893 |
| TTO—D2     | -0.027846         | 0.034067 | -0.817387 | 0.4137 |
| 340 TUN—D2 | 0.009487          | 0.102131 | 0.092890  | 0.9260 |
| TUR—D2     | C(341)= 0.020144  | 0.031600 | 0.637467  | 0.5238 |
| URY—D2     | 0.005043          | 0.024769 | 0.203596  | 0.8387 |
| USA—D2     | C(343)= 0.014198  | 0.064549 | 0.219949  | 0.8259 |
| VCT—D2     | 0.086512          | 0.093802 | 0.922284  | 0.3564 |
| 345 ZAF—D2 | C(345)= -0.087301 | 0.036817 | -2.371223 | 0.0178 |
| ZMB—D2     | -0.132269         | 0.131414 | -1.006509 | 0.3142 |
| ZWE—D2     | -0.021719         | 0.126388 | -0.171845 | 0.8636 |
| AND—D3     | -0.210017         | 0.074895 | -2.804144 | 0.0051 |
| ARG—D3     | 0.006453          | 0.026224 | 0.246083  | 0.8056 |
| 350 AUS—D3 | -0.122572         | 0.054539 | -2.247417 | 0.0247 |

|     |        |                  |          |           |        |
|-----|--------|------------------|----------|-----------|--------|
|     | AUT—D3 | -0.121918        | 0.045804 | -2.661698 | 0.0078 |
|     | BDI—D3 | 0.394702         | 0.178163 | 2.215396  | 0.0268 |
|     | BEL—D3 | -0.119476        | 0.045564 | -2.622171 | 0.0088 |
|     | BEN—D3 | 0.311241         | 0.124548 | 2.498962  | 0.0125 |
| 355 | BFA—D3 | 0.431952         | 0.157490 | 2.742728  | 0.0061 |
|     | BGD—D3 | 0.467028         | 0.158783 | 2.941303  | 0.0033 |
|     | BHS—D3 | -0.115470        | 0.026404 | -4.373148 | 0.0000 |
|     | BLZ—D3 | 0.224827         | 0.078220 | 2.874309  | 0.0041 |
|     | BOL—D3 | 0.190689         | 0.085514 | 2.229916  | 0.0258 |
| 360 | BRA—D3 | C(360)= 0.010909 | 0.019224 | 0.567486  | 0.5704 |
|     | BWA—D3 | 0.227827         | 0.088169 | 2.583991  | 0.0098 |
|     | CAF—D3 | 0.291505         | 0.142444 | 2.046447  | 0.0408 |
|     | CAN—D3 | -0.127135        | 0.051898 | -2.449691 | 0.0143 |
|     | CHE—D3 | -0.219456        | 0.077236 | -2.841362 | 0.0045 |
| 365 | CHL—D3 | 0.146060         | 0.049319 | 2.961537  | 0.0031 |
|     | CHN—D3 | C(366)= 0.587545 | 0.156993 | 3.973932  | 0.0001 |
|     | CIV—D3 | 0.078627         | 0.087776 | 0.895765  | 0.3704 |
|     | CMR—D3 | 0.221430         | 0.113258 | 1.955089  | 0.0506 |
|     | COD—D3 | 0.186728         | 0.128791 | 1.449855  | 0.1472 |
| 370 | COG—D3 | 0.177897         | 0.082165 | 2.165118  | 0.0304 |
|     | COL—D3 | 0.123732         | 0.050217 | 2.463962  | 0.0138 |
|     | CRI—D3 | 0.086333         | 0.036673 | 2.354116  | 0.0186 |
|     | CUB—D3 | 0.171236         | 0.060248 | 2.842204  | 0.0045 |
|     | DEU—D3 | -0.121802        | 0.041903 | -2.906779 | 0.0037 |
| 375 | DNK—D3 | -0.164285        | 0.061620 | -2.666096 | 0.0077 |
|     | DOM—D3 | 0.188311         | 0.065224 | 2.887132  | 0.0039 |
|     | DZA—D3 | 0.118609         | 0.056741 | 2.090348  | 0.0366 |
|     | ECU—D3 | 0.107011         | 0.050889 | 2.102838  | 0.0355 |
|     | EGY—D3 | 0.332689         | 0.111121 | 2.993942  | 0.0028 |
| 380 | ESP—D3 | -0.067797        | 0.032447 | -2.089472 | 0.0367 |
|     | FIN—D3 | -0.096217        | 0.044743 | -2.150454 | 0.0316 |
|     | FJI—D3 | 0.123431         | 0.067590 | 1.826160  | 0.0679 |
|     | FRA—D3 | -0.129773        | 0.043423 | -2.988619 | 0.0028 |
|     | GAB—D3 | -0.146796        | 0.067109 | -2.187447 | 0.0288 |
| 385 | GBR—D3 | -0.091325        | 0.038172 | -2.392490 | 0.0168 |
|     | GEO—D3 | 0.105432         | 0.064331 | 1.638905  | 0.1013 |
|     | GHA—D3 | 0.300313         | 0.119899 | 2.504712  | 0.0123 |
|     | GMB—D3 | 0.269352         | 0.118615 | 2.270816  | 0.0232 |
|     | GNB—D3 | 0.321053         | 0.144526 | 2.221420  | 0.0264 |

|     |        |                  |          |           |        |
|-----|--------|------------------|----------|-----------|--------|
| 390 | GRC—D3 | -0.082832        | 0.032415 | -2.555396 | 0.0106 |
|     | GRL—D3 | -0.101271        | 0.037024 | -2.735290 | 0.0063 |
|     | GTM—D3 | 0.140218         | 0.071890 | 1.950452  | 0.0512 |
|     | GUY—D3 | 0.192648         | 0.073219 | 2.631141  | 0.0085 |
|     | HKG—D3 | 0.061659         | 0.024679 | 2.498446  | 0.0125 |
| 395 | HND—D3 | 0.215299         | 0.093338 | 2.306667  | 0.0211 |
|     | HTI—D3 | 0.189798         | 0.122924 | 1.544030  | 0.1226 |
|     | IDN—D3 | 0.318226         | 0.105666 | 3.011635  | 0.0026 |
|     | IND—D3 | 0.466032         | 0.137864 | 3.380383  | 0.0007 |
|     | IRL—D3 | -0.019808        | 0.038315 | -0.516968 | 0.6052 |
| 400 | IRN—D3 | -0.020130        | 0.043362 | -0.464225 | 0.6425 |
|     | IRQ—D3 | 0.165432         | 0.109843 | 1.506082  | 0.1321 |
|     | ISL—D3 | -0.110947        | 0.039910 | -2.779933 | 0.0055 |
|     | ISR—D3 | -0.082211        | 0.029823 | -2.756677 | 0.0059 |
|     | ITA—D3 | -0.114244        | 0.037851 | -3.018255 | 0.0026 |
| 405 | JAM—D3 | 0.082406         | 0.044627 | 1.846572  | 0.0649 |
|     | JPN—D3 | -0.116628        | 0.044075 | -2.646105 | 0.0082 |
|     | KEN—D3 | 0.253720         | 0.131662 | 1.927051  | 0.0540 |
|     | KIR—D3 | 0.049922         | 0.078568 | 0.635399  | 0.5252 |
|     | KOR—D3 | 0.228768         | 0.055749 | 4.103526  | 0.0000 |
| 410 | LKA—D3 | 0.371652         | 0.111856 | 3.322595  | 0.0009 |
|     | LSO—D3 | 0.389750         | 0.156061 | 2.497426  | 0.0125 |
|     | LUX—D3 | -0.150247        | 0.074134 | -2.026701 | 0.0427 |
|     | MAR—D3 | 0.283408         | 0.096003 | 2.952067  | 0.0032 |
|     | MCO—D3 | -0.309067        | 0.112807 | -2.739787 | 0.0062 |
| 415 | MDG—D3 | 0.266059         | 0.136459 | 1.949734  | 0.0513 |
|     | MEX—D3 | C(416)= 0.013689 | 0.021505 | 0.636532  | 0.5245 |
|     | MLI—D3 | 0.388327         | 0.154156 | 2.519057  | 0.0118 |
|     | MLT—D3 | 0.059287         | 0.033865 | 1.750706  | 0.0801 |
|     | MMR—D3 | 0.632622         | 0.184431 | 3.430127  | 0.0006 |
| 420 | MRT—D3 | 0.177290         | 0.085665 | 2.069577  | 0.0385 |
|     | MWI—D3 | 0.363313         | 0.162467 | 2.236222  | 0.0254 |
|     | MYS—D3 | C(422)= 0.185691 | 0.058642 | 3.166513  | 0.0016 |
|     | NER—D3 | 0.277726         | 0.151222 | 1.836544  | 0.0663 |
|     | NGA—D3 | 0.181795         | 0.086790 | 2.094657  | 0.0363 |
| 425 | NIC—D3 | 0.155846         | 0.070975 | 2.195782  | 0.0282 |
|     | NLD—D3 | -0.136292        | 0.053738 | -2.536207 | 0.0112 |
|     | NOR—D3 | -0.181870        | 0.073424 | -2.476984 | 0.0133 |
|     | NPL—D3 | 0.467433         | 0.165540 | 2.823689  | 0.0048 |



|            |                   |          |           |        |
|------------|-------------------|----------|-----------|--------|
| NZL—D3     | -0.103698         | 0.041599 | -2.492770 | 0.0127 |
| 430 OMN—D3 | 0.015879          | 0.039162 | 0.405482  | 0.6851 |
| PAK—D3     | 0.394437          | 0.137393 | 2.870861  | 0.0041 |
| PAN—D3     | 0.136059          | 0.047834 | 2.844412  | 0.0045 |
| PER—D3     | 0.122595          | 0.048204 | 2.543276  | 0.0110 |
| PHL—D3     | 0.210326          | 0.089503 | 2.349920  | 0.0188 |
| 435 PNG—D3 | 0.183287          | 0.092265 | 1.986528  | 0.0470 |
| PRI—D3     | -0.033915         | 0.014962 | -2.266731 | 0.0234 |
| PRT—D3     | -0.023073         | 0.011636 | -1.982942 | 0.0474 |
| PRY—D3     | 0.145025          | 0.074300 | 1.951885  | 0.0510 |
| RWA—D3     | 0.429392          | 0.167543 | 2.562872  | 0.0104 |
| 440 SAU—D3 | -0.275597         | 0.059478 | -4.633626 | 0.0000 |
| SDN—D3     | 0.335161          | 0.104255 | 3.214830  | 0.0013 |
| SEN—D3     | 0.244358          | 0.101687 | 2.403037  | 0.0163 |
| SGP—D3     | 0.029459          | 0.029684 | 0.992422  | 0.3210 |
| SLE—D3     | 0.348502          | 0.147604 | 2.361068  | 0.0183 |
| 445 SLV—D3 | 0.113740          | 0.059638 | 1.907180  | 0.0566 |
| SUR—D3     | 0.026326          | 0.031632 | 0.832263  | 0.4053 |
| SWE—D3     | -0.134449         | 0.058527 | -2.297215 | 0.0216 |
| SWZ—D3     | 0.263466          | 0.077681 | 3.391626  | 0.0007 |
| SYC—D3     | 0.055916          | 0.053481 | 1.045530  | 0.2958 |
| 450 TCD—D3 | 0.412879          | 0.118398 | 3.487199  | 0.0005 |
| TGO—D3     | 0.274618          | 0.133373 | 2.059027  | 0.0395 |
| THA—D3     | 0.324162          | 0.093926 | 3.451232  | 0.0006 |
| TTO—D3     | 0.092230          | 0.021408 | 4.308275  | 0.0000 |
| TUN—D3     | 0.227384          | 0.082235 | 2.765037  | 0.0057 |
| 455 TUR—D3 | C(455)= 0.094700  | 0.019641 | 4.821570  | 0.0000 |
| URY—D3     | 0.050878          | 0.030237 | 1.682650  | 0.0925 |
| USA—D3     | C(457)= -0.129767 | 0.050865 | -2.551186 | 0.0108 |
| VCT—D3     | 0.244883          | 0.083653 | 2.927362  | 0.0034 |
| ZAF—D3     | C(459)= 0.035947  | 0.025307 | 1.420404  | 0.1556 |
| 460 ZMB—D3 | 0.234432          | 0.096229 | 2.436200  | 0.0149 |
| ZWE—D3     | 0.151329          | 0.117630 | 1.286486  | 0.1983 |
| AND—D4     | -0.147310         | 0.072895 | -2.020852 | 0.0433 |
| ARG—D4     | -0.040656         | 0.024639 | -1.650042 | 0.0990 |
| AUS—D4     | -0.039063         | 0.054694 | -0.714197 | 0.4751 |
| 465 AUT—D4 | -0.046465         | 0.044217 | -1.050838 | 0.2934 |
| BDI—D4     | 0.027158          | 0.175400 | 0.154836  | 0.8770 |
| BEL—D4     | -0.049542         | 0.044638 | -1.109869 | 0.2671 |

|     |        |                   |          |           |        |
|-----|--------|-------------------|----------|-----------|--------|
|     | BEN—D4 | 0.092423          | 0.121641 | 0.759804  | 0.4474 |
|     | BFA—D4 | 0.187626          | 0.154476 | 1.214598  | 0.2246 |
| 470 | BGD—D4 | 0.267492          | 0.152436 | 1.754782  | 0.0794 |
|     | BHS—D4 | -0.090739         | 0.023403 | -3.877184 | 0.0001 |
|     | BLZ—D4 | 0.082265          | 0.078139 | 1.052808  | 0.2925 |
|     | BOL—D4 | 0.045303          | 0.078713 | 0.575550  | 0.5649 |
|     | BRA—D4 | C(474)= -0.036503 | 0.021156 | -1.725425 | 0.0845 |
| 475 | BWA—D4 | 0.179471          | 0.088764 | 2.021888  | 0.0432 |
|     | CAF—D4 | -0.053309         | 0.151853 | -0.351056 | 0.7256 |
|     | CAN—D4 | -0.052678         | 0.050808 | -1.036805 | 0.2999 |
|     | CHE—D4 | -0.106648         | 0.076302 | -1.397716 | 0.1623 |
|     | CHL—D4 | 0.129442          | 0.049067 | 2.638047  | 0.0084 |
| 480 | CHN—D4 | C(480)= 0.562153  | 0.151754 | 3.704379  | 0.0002 |
|     | CIV—D4 | -0.081399         | 0.081317 | -1.001010 | 0.3169 |
|     | CMR—D4 | 0.021195          | 0.109536 | 0.193501  | 0.8466 |
|     | COD—D4 | -0.091254         | 0.124095 | -0.735356 | 0.4622 |
|     | COG—D4 | 0.014735          | 0.077938 | 0.189056  | 0.8501 |
| 485 | COL—D4 | 0.062156          | 0.046966 | 1.323425  | 0.1858 |
|     | CRI—D4 | 0.039695          | 0.033724 | 1.177058  | 0.2392 |
|     | CUB—D4 | 0.072118          | 0.054318 | 1.327691  | 0.1843 |
|     | DEU—D4 | -0.036534         | 0.041687 | -0.876380 | 0.3809 |
|     | DNK—D4 | -0.063914         | 0.059195 | -1.079736 | 0.2803 |
| 490 | DOM—D4 | 0.135496          | 0.061554 | 2.201257  | 0.0278 |
|     | DZA—D4 | -0.004317         | 0.054681 | -0.078951 | 0.9371 |
|     | ECU—D4 | 0.003716          | 0.046699 | 0.079574  | 0.9366 |
|     | EGY—D4 | 0.173843          | 0.108373 | 1.604118  | 0.1088 |
|     | ESP—D4 | -0.034271         | 0.031850 | -1.076025 | 0.2820 |
| 495 | FIN—D4 | -0.029957         | 0.040975 | -0.731099 | 0.4648 |
|     | FJI—D4 | 0.030932          | 0.064619 | 0.478682  | 0.6322 |
|     | FRA—D4 | -0.062629         | 0.042323 | -1.479814 | 0.1390 |
|     | GAB—D4 | -0.183949         | 0.066919 | -2.748847 | 0.0060 |
|     | GBR—D4 | -0.027236         | 0.036921 | -0.737693 | 0.4607 |
| 500 | GEO—D4 | -0.001072         | 0.059480 | -0.018023 | 0.9856 |
|     | GHA—D4 | 0.136971          | 0.114981 | 1.191253  | 0.2336 |
|     | GMB—D4 | -0.000908         | 0.116872 | -0.007769 | 0.9938 |
|     | GNB—D4 | 0.052652          | 0.141142 | 0.373043  | 0.7091 |
|     | GRC—D4 | -0.114642         | 0.030334 | -3.779343 | 0.0002 |
| 505 | GRL—D4 | -0.042445         | 0.037134 | -1.143019 | 0.2531 |
|     | GTM—D4 | 0.005903          | 0.069150 | 0.085368  | 0.9320 |

|     |        |                   |          |           |        |
|-----|--------|-------------------|----------|-----------|--------|
|     | GUY—D4 | 0.090652          | 0.064930 | 1.396149  | 0.1627 |
|     | HKG—D4 | 0.115658          | 0.024916 | 4.641941  | 0.0000 |
|     | HND—D4 | 0.040926          | 0.092269 | 0.443555  | 0.6574 |
| 510 | HTI—D4 | -0.067303         | 0.117201 | -0.574253 | 0.5658 |
|     | IDN—D4 | 0.212224          | 0.101316 | 2.094682  | 0.0362 |
|     | IND—D4 | 0.307837          | 0.131189 | 2.346516  | 0.0190 |
|     | IRL—D4 | 0.110926          | 0.045535 | 2.436081  | 0.0149 |
|     | IRN—D4 | -0.120718         | 0.047688 | -2.531399 | 0.0114 |
| 515 | IRQ—D4 | 0.097424          | 0.083332 | 1.169114  | 0.2424 |
|     | ISL—D4 | -0.038892         | 0.041294 | -0.941837 | 0.3463 |
|     | ISR—D4 | -0.023983         | 0.029259 | -0.819682 | 0.4124 |
|     | ITA—D4 | -0.072506         | 0.035119 | -2.064592 | 0.0390 |
|     | JAM—D4 | -0.035739         | 0.046348 | -0.771106 | 0.4407 |
| 520 | JPN—D4 | -0.032880         | 0.042233 | -0.778540 | 0.4363 |
|     | KEN—D4 | 0.054845          | 0.128016 | 0.428420  | 0.6684 |
|     | KIR—D4 | -0.133806         | 0.075006 | -1.783938 | 0.0745 |
|     | KOR—D4 | 0.256889          | 0.055965 | 4.590157  | 0.0000 |
|     | LKA—D4 | 0.264714          | 0.107917 | 2.452949  | 0.0142 |
| 525 | LSO—D4 | 0.176526          | 0.149849 | 1.178029  | 0.2388 |
|     | LUX—D4 | -0.015538         | 0.071388 | -0.217654 | 0.8277 |
|     | MAR—D4 | 0.140099          | 0.090081 | 1.555260  | 0.1199 |
|     | MCO—D4 | -0.106593         | 0.112712 | -0.945704 | 0.3443 |
|     | MDG—D4 | -0.032366         | 0.133977 | -0.241576 | 0.8091 |
| 530 | MEX—D4 | C(530)= -0.024957 | 0.022820 | -1.093623 | 0.2742 |
|     | MLI—D4 | 0.122437          | 0.148316 | 0.825514  | 0.4091 |
|     | MLT—D4 | 0.107646          | 0.035030 | 3.072975  | 0.0021 |
|     | MMR—D4 | 0.415968          | 0.175133 | 2.375158  | 0.0176 |
|     | MRT—D4 | -0.011334         | 0.085867 | -0.131995 | 0.8950 |
| 535 | MWI—D4 | 0.079309          | 0.159692 | 0.496640  | 0.6195 |
|     | MYS—D4 | C(536)= 0.165387  | 0.058394 | 2.832254  | 0.0046 |
|     | NER—D4 | 0.009084          | 0.146893 | 0.061838  | 0.9507 |
|     | NGA—D4 | -0.002931         | 0.081156 | -0.036111 | 0.9712 |
|     | NIC—D4 | -0.021905         | 0.060486 | -0.362156 | 0.7173 |
| 540 | NLD—D4 | -0.054501         | 0.052982 | -1.028667 | 0.3037 |
|     | NOR—D4 | -0.061501         | 0.072520 | -0.848056 | 0.3964 |
|     | NPL—D4 | 0.233339          | 0.159262 | 1.465124  | 0.1429 |
|     | NZL—D4 | -0.047491         | 0.041742 | -1.137740 | 0.2553 |
|     | OMN—D4 | -0.031444         | 0.039054 | -0.805142 | 0.4208 |
| 545 | PAK—D4 | 0.166839          | 0.134176 | 1.243430  | 0.2138 |
|     | PAN—D4 | 0.121357          | 0.042383 | 2.863338  | 0.0042 |
|     | PER—D4 | 0.046795          | 0.041968 | 1.115032  | 0.2649 |

|            |                   |          |           |        |
|------------|-------------------|----------|-----------|--------|
| PHL—D4     | 0.090230          | 0.085608 | 1.053992  | 0.2919 |
| PNG—D4     | 0.049766          | 0.085765 | 0.580263  | 0.5618 |
| 550 PRI—D4 | -0.012215         | 0.013937 | -0.876472 | 0.3808 |
| PRT—D4     | -0.006921         | 0.011675 | -0.592785 | 0.5534 |
| PRY—D4     | 0.070363          | 0.072466 | 0.970973  | 0.3316 |
| RWA—D4     | 0.185095          | 0.160626 | 1.152335  | 0.2492 |
| SAU—D4     | -0.261954         | 0.058306 | -4.492744 | 0.0000 |
| 555 SDN—D4 | 0.141678          | 0.093303 | 1.518478  | 0.1290 |
| SEN—D4     | 0.045443          | 0.098935 | 0.459322  | 0.6460 |
| SGP—D4     | 0.135323          | 0.023284 | 5.811984  | 0.0000 |
| SLE—D4     | 0.059652          | 0.153779 | 0.387905  | 0.6981 |
| SLV—D4     | -0.012744         | 0.057274 | -0.222514 | 0.8239 |
| 560 SUR—D4 | -0.059037         | 0.031508 | -1.873688 | 0.0610 |
| SWE—D4     | -0.044557         | 0.057087 | -0.780511 | 0.4351 |
| SWZ—D4     | 0.151934          | 0.075574 | 2.010394  | 0.0444 |
| SYC—D4     | 0.064994          | 0.050926 | 1.276244  | 0.2019 |
| TCD—D4     | 0.132028          | 0.116045 | 1.137723  | 0.2553 |
| 565 TGO—D4 | 0.039921          | 0.128942 | 0.309606  | 0.7569 |
| THA—D4     | 0.240192          | 0.093272 | 2.575177  | 0.0100 |
| TTO—D4     | 0.029883          | 0.022267 | 1.342027  | 0.1796 |
| TUN—D4     | 0.094006          | 0.079184 | 1.187185  | 0.2352 |
| TUR—D4     | C(569)= 0.100434  | 0.019974 | 5.028275  | 0.0000 |
| 570 URY—D4 | 0.053515          | 0.021526 | 2.486007  | 0.0130 |
| USA—D4     | C(571)= -0.040304 | 0.048932 | -0.823685 | 0.4102 |
| VCT—D4     | 0.138755          | 0.083205 | 1.667634  | 0.0955 |
| ZAF—D4     | C(573)= -0.053242 | 0.024067 | -2.212231 | 0.0270 |
| ZMB—D4     | 0.035484          | 0.088075 | 0.402878  | 0.6871 |
| 575 ZWE—D4 | 0.024157          | 0.103241 | 0.233989  | 0.8150 |

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Weighted Statistics

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|                    |          |                    |           |
|--------------------|----------|--------------------|-----------|
| R-squared          | 0.342232 | Mean dependent var | -0.000358 |
| Adjusted R-squared | 0.266886 | S.D. dependent var | 0.053221  |
| S.E. of regression | 0.045569 | Sum squared resid  | 10.40535  |
| F-statistic        | 4.542147 | Durbin-Watson stat | 1.706694  |
| Prob(F-statistic)  | 0.000000 |                    |           |

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Unweighted Statistics

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|                   |          |                    |           |
|-------------------|----------|--------------------|-----------|
| R-squared         | 0.290712 | Mean dependent var | -0.001243 |
| Sum squared resid | 10.52293 | Durbin-Watson stat | 1.899193  |

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## Appendix C

### The Results of Wald Tests (Outputs of Eviews)

1. The Result of the Wald Test for  $H_0 : \beta_1 = 0$

Wald Test:

| <i>Test Statistic</i> | <i>Value</i> | <i>df</i> | <i>Probability</i> |
|-----------------------|--------------|-----------|--------------------|
| F-statistic           | 18.57513     | (1, 5011) | 0.0000             |
| Chi-square            | 18.57513     | 1         | 0.0000             |

Null Hypothesis Summary:

| Normalized Restriction (= 0) | Value    | Std. Err. |
|------------------------------|----------|-----------|
| -C(1) - C(2)                 | 0.120412 | 0.027938  |

Restrictions are linear in coefficients.

Note:  $-C(1) = \hat{\beta}_0$  ;  $C(2) = \hat{\lambda}_1$

2. The Result of the Wald Test for  $:= 0$

Wald Test:

| <i>Test Statistic</i> | <i>Value</i> | <i>df</i> | <i>Probability</i> |
|-----------------------|--------------|-----------|--------------------|
| F-statistic           | 20.05227     | (1, 5011) | 0.0000             |
| Chi-square            | 20.05227     | 1         | 0.0000             |

Null Hypothesis Summary:

| Normalized Restriction (= 0) | Value    | Std. Err. |
|------------------------------|----------|-----------|
| -C(1) - C(3)                 | 0.192575 | 0.043005  |

Restrictions are linear in coefficients.

Note:  $-C(1) = \hat{\beta}_0$  ;  $C(3) = \hat{\lambda}_1$

3. The Result of the Wald Test for  $: = 0$

Wald Test:

| <i>Test Statistic</i> | <i>Value</i> | <i>df</i> | <i>Probability</i> |
|-----------------------|--------------|-----------|--------------------|
| F-statistic           | 7.966911     | (1, 5011) | 0.0048             |
| Chi-square            | 7.966911     | 1         | 0.0048             |

Null Hypothesis Summary:

| Normalized Restriction (= 0) | Value    | Std. Err. |
|------------------------------|----------|-----------|
| -C(1) - C(4)                 | 0.068187 | 0.024158  |

Restrictions are linear in coefficients.

Note:  $-C(1) = \hat{\beta}_0$  ;  $C(4) = \hat{\lambda}_3$

4. The Result of the Wald Test for : = 0

Wald Test:

| Test Statistic | Value    | df        | Probability |
|----------------|----------|-----------|-------------|
| F-statistic    | 41.76206 | (1, 5011) | 0.0000      |
| Chi-square     | 41.76206 | 1         | 0.0000      |

Null Hypothesis Summary:

| Normalized Restriction (= 0) | Value    | Std. Err. |
|------------------------------|----------|-----------|
| -C(1) - C(5)                 | 0.149236 | 0.023093  |

Restrictions are linear in coefficients.

Note:  $-C(1) = \hat{\beta}_0$  ;  $C(5) = \hat{\lambda}_4$

5. The Result of the Wald Test for :  $y_1^*(BRA) - y_0^*(BRA) = 0$

Wald Test:

| Test Statistic | Value    | df        | Probability |
|----------------|----------|-----------|-------------|
| F-statistic    | 2.168572 | (1, 5011) | 0.1409      |
| Chi-square     | 2.168572 | 1         | 0.1409      |

Null Hypothesis Summary:

| Normalized Restriction (= 0)                   | Value     | Std. Err. |
|--|-----------|-----------|
| $C(18)/C(1) + (C(18) + C(132))/(-C(1) - C(2))$ | -0.158940 | 0.107931  |

Delta method computed using analytic derivatives.

Note:  $C(18)/C(1) = -C(18)/(-C(1)) = -\hat{c}_0(BRA)/\hat{\beta}_0$  ;  
 $(C(18) + C(132))/(-C(1) - C(2)) = \hat{c}_1(BRA)/\hat{\beta}_1$

6. The Result of the Wald Test for  $H_0 : y_2^*(BRA) - y_1^*(BRA) = 0$

Wald Test:

| <i>Test Statistic</i> | <i>Value</i> | <i>df</i> | <i>Probability</i> |
|-----------------------|--------------|-----------|--------------------|
| F-statistic           | 1.367203     | (1, 5011) | 0.2423             |
| Chi-square            | 1.367203     | 1         | 0.2423             |

Null Hypothesis Summary:

| <i>Normalized Restriction (= 0)</i>                                  | <i>Value</i> | <i>Std. Err.</i> |
|--|--------------|------------------|
| $-(C(18) + C(132))/(-C(1) - C(2)) + (C(18) + C(246))/(-C(1) - C(3))$ | -0.138193    | 0.118187         |

Delta method computed using analytic derivatives.

Note:  $(C(18) + C(132))/(-C(1) - C(2)) = \hat{c}_1(BRA) / \hat{\beta}_1;$

$(C(18) + C(246))/(-C(1) - C(3)) = (BRA) /$

7. The Result of the Wald Test for  $H_0 : y_3^*(BRA) - y_2^*(BRA) = 0$

Wald Test:

| <i>Test Statistic</i> | <i>Value</i> | <i>df</i> | <i>Probability</i> |
|-----------------------|--------------|-----------|--------------------|
| F-statistic           | 0.569852     | (1, 5011) | 0.4504             |
| Chi-square            | 0.569852     | 1         | 0.4503             |

Null Hypothesis Summary:

| <i>Normalized Restriction (= 0)</i>                                  | <i>Value</i> | <i>Std. Err.</i> |
|--|--------------|------------------|
| $-(C(18) + C(246))/(-C(1) - C(3)) + (C(18) + C(360))/(-C(1) - C(4))$ | 0.087458     | 0.115855         |

Delta method computed using analytic derivatives.

Note:  $(C(18) + C(246))/(-C(1) - C(3)) = \hat{c}_2(BRA) / \hat{\beta}_2;$

$(C(18) + C(360))/(-C(1) - C(4)) = \hat{c}_3(BRA) / \hat{\beta}_3$

8. The Result of the Wald Test for  $H_0 : y_4^* (BRA) - y_3^* (BRA) = 0$

Wald Test:

| <i>Test Statistic</i> | <i>Value</i> | <i>df</i> | <i>Probability</i> |
|-----------------------|--------------|-----------|--------------------|
| F-statistic           | 0.587786     | (1, 5011) | 0.4433             |
| Chi-square            | 0.587786     | 1         | 0.4433             |

Null Hypothesis Summary:

| <i>Normalized Restriction (= 0)</i>                                  | <i>Value</i> | <i>Std. Err.</i> |
|--|--------------|------------------|
| $-(C(18) + C(360))/(-C(1) - C(4)) + (C(18) + C(474))/(-C(1) - C(5))$ | -0.094479    | 0.123233         |

Delta method computed using analytic derivatives.

Note:  $(C(18) + C(360))/(-C(1) - C(4)) = \hat{c}_3 (BRA) / \hat{\beta}_3 ;$

$$(C(18) + C(474))/(-C(1) - C(5)) = \hat{c}_4 (BRA) / \hat{\beta}_4$$

9. The Result of the Wald Test for  $H_0 : y_0^* (BRA) = 0$

Wald Test:

| <i>Test Statistic</i> | <i>Value</i> | <i>df</i> | <i>Probability</i> |
|-----------------------|--------------|-----------|--------------------|
| F-statistic           | 65.09192     | (1, 5011) | 0.0000             |
| Chi-square            | 65.09192     | 1         | 0.0000             |

Null Hypothesis Summary:

| <i>Normalized Restriction (= 0)</i> | <i>Value</i> | <i>Std. Err.</i> |
|-------------------------------------|--------------|------------------|
| $-C(18) / C(1)$                     | -0.201337    | 0.024955         |

Delta method computed using analytic derivatives.

Note:  $-C(18)/C(1) = C(18)/(-C(1)) = \hat{c}_0 (BRA) / \hat{\beta}_0$

10. The Result of the Wald Test for  $H_0 : y_1^* (BRA) = 0$

Wald Test:



| <i>Test Statistic</i> | <i>Value</i> | <i>df</i> | <i>Probability</i> |
|-----------------------|--------------|-----------|--------------------|
| F-statistic           | 11.77179     | (1, 5011) | 0.0006             |
| Chi-square            | 11.77179     | 1         | 0.0006             |

Null Hypothesis Summary:

| <i>Normalized Restriction (= 0)</i> | <i>Value</i> | <i>Std. Err.</i> |
|-------------------------------------|--------------|------------------|
| (C(18) + C(132)) / (-C(1) - C(2))   | -0.360277    | 0.105006         |

Delta method computed using analytic derivatives.

Note:  $(C(18) + C(132))/(-C(1) - C(2)) = \hat{c}_1 (BRA) / \hat{\beta}_1$

11. The Result of the Wald Test for  $H_0: y_2^* (BRA) = 0$

Wald Test:

| <i>Test Statistic</i> | <i>Value</i> | <i>df</i> | <i>Probability</i> |
|-----------------------|--------------|-----------|--------------------|
| F-statistic           | 84.46333     | (1, 5011) | 0.0000             |
| Chi-square            | 84.46333     | 1         | 0.0000             |

Null Hypothesis Summary:

| <i>Normalized Restriction (= 0)</i> | <i>Value</i> | <i>Std. Err.</i> |
|-------------------------------------|--------------|------------------|
| (C(18) + C(246)) / (-C(1) - C(3))   | -0.498469    | 0.054238         |

Delta method computed using analytic derivatives.

Note:  $(C(18) + C(246))/(-C(1) - C(3)) = \hat{c}_2 (BRA) / \hat{\beta}_2$

12. The Result of the Wald Test for  $H_0: y_3^* (BRA) = 0$

Wald Test:

| <i>Test Statistic</i> | <i>Value</i> | <i>df</i> | <i>Probability</i> |
|-----------------------|--------------|-----------|--------------------|
| F-statistic           | 16.11826     | (1, 5011) | 0.0001             |
| Chi-square            | 16.11826     | 1         | 0.0001             |

Null Hypothesis Summary:

| <i>Normalized Restriction (= 0)</i> | <i>Value</i> | <i>Std. Err.</i> |
|-------------------------------------|--------------|------------------|
| (C(18) + C(360)) / (-C(1) - C(4))   | -0.411012    | 0.102375         |

Delta method computed using analytic derivatives.

Note:  $(C(18) + C(360)) / (-C(1) - C(4)) = \hat{c}_3 (BRA) / \hat{\beta}_3$

13. The Result of the Wald Test for  $H_0: y_4^* (BRA) = 0$

Wald Test:

| Test Statistic | Value    | df        | Probability |
|----------------|----------|-----------|-------------|
| F-statistic    | 54.30023 | (1, 5011) | 0.0000      |
| Chi-square     | 54.30023 | 1         | 0.0000      |

Null Hypothesis Summary:

| Normalized Restriction (= 0)        | Value     | Std. Err. |
|-------------------------------------|-----------|-----------|
| $(C(18) + C(474)) / (-C(1) - C(5))$ | -0.505491 | 0.068598  |

Delta method computed using analytic derivatives.

Note:  $(C(18) + C(474)) / (-C(1) - C(5)) = \hat{c}_4 (BRA) / \hat{\beta}_4$

14. The Result of the Wald Test for  $H_0: y_0^* (CHN) - y_0^* (BRA) = 0$

Wald Test:

| Test Statistic | Value    | df        | Probability |
|----------------|----------|-----------|-------------|
| F-statistic    | 1357.977 | (1, 5011) | 0.0000      |
| Chi-square     | 1357.977 | 1         | 0.0000      |

Null Hypothesis Summary:

| Normalized Restriction (= 0) | Value     | Std. Err. |
|------------------------------|-----------|-----------|
| $C(18)/C(1) - C(24)/C(1)$    | -3.279301 | 0.088989  |

Delta method computed using analytic derivatives.

Note:  $-C(24)/C(1) = C(24)/(-C(1)) = \hat{c}_0 (CHN) / \hat{\beta}_0$  ;

$C(18)/C(1) = -C(18)/(-C(1)) = -\hat{c}_0 (BRA) / \hat{\beta}_0$

15. The Result of the Wald Test for  $H_0: y_1^* (CHN) - y_1^* (BRA) = 0$

Wald Test:

| <i>Test Statistic</i> | <i>Value</i> | <i>df</i> | <i>Probability</i> |
|-----------------------|--------------|-----------|--------------------|
| F-statistic           | 229.4352     | (1, 5011) | 0.0000             |
| Chi-square            | 229.4352     | 1         | 0.0000             |

Null Hypothesis Summary:

| <i>Normalized Restriction (= 0)</i>                                  | <i>Value</i> | <i>Std. Err.</i> |
|--|--------------|------------------|
| $-(C(18) + C(132))/(-C(1) - C(2)) + (C(24) + C(138))/(-C(1) - C(2))$ | -2.249928    | 0.148538         |

Delta method computed using analytic derivatives.

Note:  $(C(24) + C(138))/(-C(1) - C(2)) = \hat{c}_1 (CHN) / \hat{\beta}_1$  ;

$$(C(18) + C(132))/(-C(1) - C(2)) = \hat{c}_1 (BRA) / \hat{\beta}_1$$

16. The Result of the Wald Test for  $H_0: y_2^* (CHN) - y_2^* (BRA) = 0$

Wald Test:

| <i>Test Statistic</i> | <i>Value</i> | <i>df</i> | <i>Probability</i> |
|-----------------------|--------------|-----------|--------------------|
| F-statistic           | 251.7001     | (1, 5011) | 0.0000             |
| Chi-square            | 251.7001     | 1         | 0.0000             |

Null Hypothesis Summary:

| <i>Normalized Restriction (= 0)</i>                                  | <i>Value</i> | <i>Std. Err.</i> |
|--|--------------|------------------|
| $-(C(18) + C(246))/(-C(1) - C(3)) + (C(24) + C(252))/(-C(1) - C(3))$ | -1.653575    | 0.104227         |

Delta method computed using analytic derivatives.

Note:  $(C(24) + C(252))/(-C(1) - C(3)) = \hat{c}_2 (CHN) / \hat{\beta}_2$  ;

$$(C(18) + C(246))/(-C(1) - C(3)) = \hat{c}_2 (BRA) / \hat{\beta}_2$$

17. The Result of the Wald Test for  $H_0: y_3^*(CHN) - y_3^*(BRA) = 0$

Wald Test:

| <i>Test Statistic</i> | <i>Value</i> | <i>df</i> | <i>Probability</i> |
|-----------------------|--------------|-----------|--------------------|
| F-statistic           | 126.0329     | (1, 5011) | 0.0000             |
| Chi-square            | 126.0329     | 1         | 0.0000             |

Null Hypothesis Summary:

| <i>Normalized Restriction (= 0)</i>                                  | <i>Value</i> | <i>Std. Err.</i> |
|--|--------------|------------------|
| $-(C(18) + C(360))/(-C(1) - C(4)) + (C(24) + C(366))/(-C(1) - C(4))$ | -0.843747    | 0.121137         |

Delta method computed using analytic derivatives.

Note:  $(C(24) + C(366))/(-C(1) - C(4)) = \hat{c}_3(CHN) / \hat{\beta}_3$  ;

$$(C(18) + C(360))/(-C(1) - C(4)) = \hat{c}_3(BRA) / \hat{\beta}_3$$

18. The Result of the Wald Test for  $H_0: y_4^*(CHN) - y_4^*(BRA) = 0$

Wald Test:

| <i>Test Statistic</i> | <i>Value</i> | <i>df</i> | <i>Probability</i> |
|-----------------------|--------------|-----------|--------------------|
| F-statistic           | 5.523785     | (1, 5011) | 0.0188             |
| Chi-square            | 5.523785     | 1         | 0.0188             |

Null Hypothesis Summary:

| <i>Normalized Restriction (= 0)</i>                                  | <i>Value</i> | <i>Std. Err.</i> |
|--|--------------|------------------|
| $-(C(18) + C(474))/(-C(1) - C(5)) + (C(24) + C(480))/(-C(1) - C(5))$ | -0.237871    | 0.101210         |

Delta method computed using analytic derivatives.

Note:  $(C(24) + C(480))/(-C(1) - C(5)) = \hat{c}_4(CHN) / \hat{\beta}_4$  ;

$$(C(18) + C(474))/(-C(1) - C(5)) = \hat{c}_4(BRA) / \hat{\beta}_4$$